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## Identification of risk factors associated with onset and progression of amyotrophic lateral sclerosis using systematic review and meta-analysis



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### ABSTRACT

Although amyotrophic lateral sclerosis (ALS) was identified as a neurological condition 150 years ago, risk factors related to the onset and progression of ALS remain largely unknown. Monogenic mutations in over 30 genes are associated with about 10% of ALS cases. The age at onset of ALS and disease types has been found to influence ALS progression. The present study was designed to identify additional putative risk factors associated with the onset and progression of ALS using systematic review and meta-analysis of observational studies. Risk factors that may be associated with ALS include: 1) genetic mutations, including the intermediate CAG repeat expansion in ATXN2; 2) previous exposure to heavy metals such as lead and mercury; 3) previous exposure to organic chemicals, such as pesticides and solvents; 4) history of electric shock; 5) history of physical trauma/injury (including head trauma/injury); 6) smoking (a weak risk factor for ALS in women); and 6) other risk factors, such as participating in professional sports, lower body mass index, lower educational attainment, or occupations requiring repetitive/strenuous work, military service, exposure to Beta-N-methylamino-L-alanin and viral infections. Risk factors that may be associated with ALS progression rate include: 1) nutritional status, including vitamin D deficiency; 2) comorbidities; 3) ethnicity and genetic factors; 4) lack of supportive care; and 4) smoking. The extent to which these associations may be causal is discussed, with further research recommended to strengthen the evidence on which determinations of causality may be based.

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## 1. Introduction

Amyotrophic lateral sclerosis (ALS) is a neurodegenerative condition affecting the voluntary motor nervous system (Sathasivam, 2010; Piaceri et al., 2011; Talman et al., 2009). The typical symptoms/signs in ALS patients include progressive atrophic weakness and spasticity of affected regions, and the spread of symptoms from the site(s) of onset. The smooth and striated

muscles of sphincters, the muscles responsible for ocular movement, and the sensory nerves (temperature, pain, taste, hearing, vision, olfaction) are usually spared. However, dementia, cognitive impairment and executive dysfunction have been detected in more than 50% of ALS patients (Phukan et al., 2007, 2011; Gordon et al., 2011b), especially the ALS patients with mutations in the *C9orf72* and *TDP43* genes (Cooper-Knock et al., 2012; Byrne et al., 2012; Ludolph et al., 2012; Seltman and Matthews, 2012). In most ALS cases, disease progression is rapid. The average survival time from diagnosis is less than 3 years, with affected individuals often dying of respiratory failure or inanition. Notably, a small proportion (~10%) of ALS cases survive for more than 10 years (Testa et al., 2004), and rare ALS cases have survived

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for more than 50 years (Tosi et al., 1994). Multiple ALS subsets and ALS variants have been described (Silani et al., 2011; Kim et al., 2009), indicating the heterogeneity of ALS phenotypes (Piaceri et al., 2011; Chio et al., 2011; Turner et al., 2013; Sabatelli et al., 2013). Different phenotypes have been reported among individuals with the same type of gene mutations, such as those of the *C9orf72* (Murray et al., 2011; Stewart et al., 2012), *SOD1* (Piaceri et al., 2011), *TDP43* and/or *FUS* genes (Mackenzie et al., 2011).

ALS is usually diagnosed within a year of the appearance of the first clinical signs (Mitchell et al., 2010). There is currently no specific diagnostic test for ALS. Instead, diagnosis relies upon consideration of differential diagnoses and the presence of unique clinical signs – upper and lower motor neuron findings confined to spinal and bulbar voluntary muscles. Pathological degeneration signs include lateral sclerosis of descending motor tracts in motor neuron cells in the brainstem and spinal anterior horns. Clinical diagnostic criteria for ALS, referred to as the EI Escorial criteria, were proposed by the World Neurology Federation in 1994, and revised in 2000 and 2008 (Silani et al., 2011). According to these criteria, ALS cases are classified as definite, probable, or possible. The World Health Organization (WHO) categorized ALS as a major subset of motor neuron disease (MND), and all subsets of MND share an identical International Classification of Diseases (ICD) code. Therefore, in the biomedical literature, MND and ALS are often used interchangeably. Misdiagnosis or misclassification has been estimated to account for about 10% of total ALS cases (Silani

ALS is a degenerative motor neuron disease that is terminal and irreversible. Disease severity can be measured using the ALS functional rating scales (ALSFRS), having a score ranging from 0–40, based on 10 basic functions in 4 key functional domains (swallowing, walking/self-care, communicating and breathing) (Anonymous, 1996). The disease course of ALS has recently been divided into 6 stages, progressing from stage 0 for no loss of function to stage 5 for death (Chio et al., 2013a,b). The progression to death of ALS is apparently slower for spinal ALS than for bulbar onset cases, with the age at onset being negatively associated with survival time (Chio et al., 2009a,b,c).

The crude incidence of ALS in Europe and North America, especially among whites, is quite stable, about 2–3 cases per 100,000 people per year (Cronin et al., 2007; Alonso et al., 2009). Although the incidence of ALS in Asia is generally lower, the highest incidences (50–100 fold greater than the global average) have been reported in Asia, specifically in Guam and the surrounding Pacific islands during the 1940–1960s (Plato et al., 2003; Okamoto et al., 2009b; Spencer et al., 2005; Waring et al., 2004; Steele, 2005; Yoshida et al., 1998). Populations of Hispanic or African origin display lower incidences of ALS (Cronin et al., 2007). Because of the short survival time of ALS patients, the mortality rate of ALS roughly reflects its incidence rate (Marin et al., 2011a).

Both the incidence and prevalence of ALS are greater in men than in women, the ratio being approximately 1.5:1 (McCombe and Henderson, 2010). Male sex is also associated with age of onset and ALS disease type (Gordon et al., 2011a). A preponderance of ALS in males has been reported for respiratory, flail arm, classic and pure lower motor neuron (*p*-LMN) phenotypes (Chio et al., 2011). This male preponderance has declined in recent decades, suggesting that changing exposure to environmental risk factors for ALS may explain part of this sex difference (Gordon et al., 2011a).

ALS is an age-associated neurodegenerative condition, having some similarities to chronic diseases such as Alzheimer's disease; however, unlike Alzheimer's disease, ALS is not strictly an aging-associated condition (Brody and Grant, 2001). The incidence of ALS increases with age, peaks at about 70 years of age, and then declines rapidly thereafter (Logroscino et al., 2008). Although ALS is a rare neurodegenerative condition, it has been reported that the

proportion of overall mortality due to ALS has increased in recent decades, more so in women than in men (Gordon et al., 2011a).

ALS cases have been classified as belonging to one of two categories - familial or sporadic, although this classification system is now being challenged. About 10% of sporadic ALS cases are caused by monogenic mutations (Chio et al., 2012). Unlike many other chronic diseases, such as various cancers, sporadic ALS cases are not randomly distributed in the general population. Geographical (Doi et al., 2010: Turabelidze et al., 2008: Malaspina et al., 2002; Ceroni et al., 1999; Poloni et al., 1997; Gunnarsson et al., 1996; Uccelli et al., 2007; Noonan et al., 2005), occupational (Gunnarsson et al., 1996; Barth et al., 2009; Rose and Brix, 2006), and conjugal (Poloni et al., 1997) ALS clusters have been observed, suggesting that shared environmental risk factors also play an important role in the etiology of ALS. Overall, genetic factors and environmental risk factors are considered to have equally important roles in the onset of ALS (Wingo et al., 2011; Al-Chalabi et al., 2010). ALS is one of most heterogeneous neurodegenerative conditions, in terms of etiology, phenotype and progression. Although many genetic and environmental risk factors have been associated with the onset of sporadic ALS in many observational studies, no single gene variant or environmental risk factor has been conclusively linked to the onset of sporadic ALS. To date, few reviews have assessed the risk factors associated with the progression of ALS. It is important for the management of ALS cases that appropriate diagnostic criteria be defined, and that known and unknown risk factors be characterized and described. The objectives of this study were to identify environmental and genetic risk factors that are associated with the onset or progression of ALS through systematic reviews and meta-analyses of observational studies.

### 2. Methods

### 2.1. Literature search

A literature search strategy was designed using a set of search terms for the disease itself, risk factors for the disease, and related terms (Supplementary material I). The search criteria were refined and optimized using PubMed (Hersi et al., 2016), and then used to search the scientific literature in other databases. The strategy focused on the searches of systematic reviews/meta-analyses and observational studies for risk factors associated with ALS disease occurrence and progression (Supplementary materials II–IV). Other databases searched included Medline, Embase, the HuGE Literature Finder (https://phgkb.cdc.gov/HuGENavigator/startPagePubLit.do), Toxiline, and Psyinfo. All searches were initially conducted through to December 2013, with an update covering the period December 2013 to February 16, 2016.

## 2.2. Literature screening

Relevant articles were identified and exported into Distiller (DistillerSR, Evidence Partners, Ottawa, Canada) software for further evaluation and review. Additional articles from manual searches were added to this database when appropriate. The articles were screened based on title and abstract, using the inclusion criteria (disease terms, study types, human study, and article types) described in Supplementary materials I–IV. The full text of the included articles was further examined for relevance according to priori criteria focusing on disease terms, study types, population information, statistical methods, results, and conclusions (Supplementary materials I–IV). Relevant data from all articles retained following full-text screening were extracted in

Distiller using a pre-designed data collection algorithm. A random sample of 5% screened articles at all screening levels was cross-examined by second reviewer in order to validate the process of article screening and data extraction.

## 2.3. Literature assessment and meta-analyses

The identified systematic reviews were assessed using the AMSTAR criteria (Shea et al., 2007). If these reviews were of moderate or good quality, they were updated with new observational studies; otherwise the review was done ab initio. The observational studies were assessed according to previously developed criteria (Turner et al., 2011; Stroup et al., 2000). If more than 3 observational articles for a given risk factor were found, then a meta-analysis of those studies was conducted with RevMan5.1 software (http://tech.cochrane.org/revman/download). The odds ratio (OR) for ALS risk due to exposure to individual environmental risk factors was used as the summary measure of risk. The primary analysis was based on random effects (RE) modelling, with a fixed effects (FE) model run as a secondary analysis. Heterogeneity across included studies was estimated by Tau<sup>2</sup>, Chi<sup>2</sup> and I<sup>2</sup> (Higgins and Thompson, 2002). Forest plots and relevant supporting statistics were also examined. Funnel plots (Begg's test) were used to evaluate possible publication bias including small study effects. Meta-analyses for subgroups, such as disease subtype, sex, and article quality, were considered if at least three articles were available in a category. ALS associated genetic factors were summarized from the existing ALS associated gene website (Lill et al., 2011), and from searches of PubMed and other databases. If the number of included observational studies is less than three, the information from these studies was described individually.

#### 3. Results

## 3.1. Risk factors associated with onset of ALS

3.1.1. Risk factors identified in available systematic reviews (SR) and meta-analyses (MA)

Two SR/MA articles had moderate quality AMSTAR scores: one SR/MA assessed the association between exposure to chemicals and ALS (Sutedia et al., 2009a,b), the other between occupation and ALS (Sutedja et al., 2007). Neither SR identified any clear association between these risk factors and ALS. One comprehensive SR/MA suggested that smoking was not a strong risk factor for ALS (Alonso et al., 2010a,b), was re-analysed with all identified observational studies (see below). Two recent SR/MAs found that previous exposure to pesticides was significantly associated with an increased risk of developing ALS (Kamel et al., 2012; Malek et al., 2012), were also re-analysed with additional observational studies. A SR/MA of observational studies examined the association between exposure to electromagnetic fields and ALS (Zhou et al., 2012). As this association may have been confounded by other related risk factors such as by occupation and electronic shock, data synthesis and discussion in depth have been provided in present study. Seven new SR/MA articles were identified in the update (Supplementary material V) and are discussed below. There appear to be unrealistic expectations that SR/MA can establish or disprove a causal relationship between risk factors and diseases (Wang and Little, 2016), and most SR/MAs may have not included sufficient information to draw firm conclusions (Belbasis et al., 2016).

## 3.1.2. Risk factors identified in primary observational studies

In the first round search, a total of 88 observational studies related to environmental risk factors was identified. In the updated

	Exposed to heavy m	etals	Conti	rol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.1.1 Exposure to Lead							
Campbell 1970	23	74	16	74	8.4%	1.63 [0.78, 3.43]	+-
Chancellor 1993	19	103	5	103	4.4%	4.43 [1.59, 12.38]	<del></del>
Deapen 1986	30	518	27	518	16.1%	1.12 [0.65, 1.91]	+
Gresham 1986	14	56	12	66	6.1%	1.50 [0.63, 3.58]	+
Kamel 2002	35	102	53	246	17.7%	1.90 [1.14, 3.17]	<b></b>
McGuire 1997	21	174	24	348	12.1%	1.85 [1.00, 3.43]	-
Pierce-Ruhland 1981	20	80	11	78	6.9%	2.03 [0.90, 4.58]	<del>  -</del>
Subtotal (95% CI)		1107		1433	71.6%	1.72 [1.33, 2.23]	◆
Total events	162		148				
Heterogeneity: Tau <sup>z</sup> = 0.	00; Chi² = 6.24, df = 6	(P = 0.4)	$0); I^2 = 49$	%			
Test for overall effect: Z =	= 4.09 (P < 0.0001)						
1.1.2 Exposure to heavy	ı matala						
		0.7	-	405	4.400	0.00 (0.04. 7.04)	<u> </u>
Binazzi 2009	9	97	7	185	4.4%	2.60 [0.94, 7.21]	
Moraham 2006	55	178	43	189	21.2%	1.52 [0.95, 2.42]	<del></del>
Provinciali 1990	7	77 <b>352</b>	4	80	2.8%	1.90 [0.53, 6.77]	
Subtotal (95% CI)		352		454	28.4%	1.69 [1.13, 2.52]	_
Total events	71	/D 00	54				
Heterogeneity: Tau <sup>2</sup> = 0.		(P = 0.6	3);	%			
Test for overall effect: Z=	= 2.55 (P = 0.01)						
Total (95% CI)		1459		1887	100.0%	1.71 [1.38, 2.11]	•
Total events	233		202				
Heterogeneity: Tau <sup>2</sup> = 0.	00; Chi <sup>2</sup> = 7.16, df= 9	(P = 0.6)	2); $I^2 = 0$ 9	%			
Test for overall effect: Z=		•					0.01 0.1 1 10 100 Favours controls Favours risk factor
Test for subgroup differe	ences: Chi² = 0.01, df	= 1 (P =	$0.94$ ), $I^2 =$	- 0%			Favours controls Favours fisk factor

**Fig. 1.** Previous exposure to lead/heavy metals is associated with increased risk of developing ALS. Exposure data were extracted from ten case-control studies, and the association between ALS and previous exposure to lead against controls was synthesized by meta-analysis using random effects model. No evidence of heterogeneity across included studies was observed, nor was there evidence of significant publication bias (data not shown).

search, 45 additional observational studies were identified (Supplementary material V). This updated information is included in the results and discussion section of the present paper.

3.1.2.1. Exposure to heavy metals. Individuals with an occupational history of exposure to heavy metals, particularly lead, were at an increased risk for ALS (Supplementary material VI). Based on seven case-control studies, meta-analysis showed that the odds ratio of having a history of exposure to lead was significantly increased for ALS patients when compared to controls [RE model, OR = 1.72 (1.33, 2.23)] (Fig. 1). There was little evidence of heterogeneity across studies ( $I^2 = 20\%$ ), or of small study effects (funnel plot, not shown).

The second most common heavy metal putatively associated with ALS is mercury. There have been case reports of ALS or ALS-like symptoms associated with exposure to mercury (Callaghan et al., 2011; Johnson and Atchison, 2009; Praline et al., 2007; Zhou et al., 2014), although the epidemiological investigation for this association has not been extensively conducted (Gresham et al., 1986)

Previous exposure to other heavy metals, such as selenium, magnesium, aluminum, manganese and cadmium, may be also associated with an increased risk of developing ALS (Ahmed and Wicklund, 2011; Vinceti et al., 2012; Cannon and Timothy Greenamyre, 2011).

3.1.2.2. Agricultural chemicals. Previous exposure to agricultural chemicals, especially to pesticides, appears to be associated with ALS (Supplementary material VII). Meta-analysis showed that the occupation of farming was associated with ALS [RE, OR = 2.18(1.23-3.84)], but with significant heterogeneity across the 10 observational studies included ( $I^2 = 76\%$ , data not shown).

However, living in a rural area or having an occupation in the agricultural sector may not be associated with onset of ALS. This association might be confounded by exposure to agricultural chemicals. A meta-analysis of 12 case control studies revealed that the OR for ALS in individuals with a history of exposure to pesticides, including agricultural chemicals, was increased by about 50% and showed moderate heterogeneity in these studies [RE model, OR = 1.57(1.25–1.98), I<sup>2</sup> = 42% (Fig. 2)], consistent with three meta-analyses (Kamel et al., 2012; Malek et al., 2012; Kang et al., 2014).

The blood levels of two pesticides were significantly increased in ALS compared to non-ALS individuals (pentachlorobenzene: OR = 2.21; 95% CI, 1.06–4.60; p = 0.04; and cis-chlordane: OR = 5.74; 95% CI, 1.80–18.20; p = 0.005) (Su et al., 2016).

3.1.2.3. Organic solvents. Previous exposure to organic solvents has been reported to be a risk factor for ALS (Table 1) (McGuire et al., 1997; Chancellor et al., 1993; Fang et al., 2009; Morahan and Pamphlett, 2006; Park et al., 2005). However, these studies did not specify what solvents were involved. Thus, articles that used the term 'solvent' were selected for meta-analysis. The OR of having ALS in individuals exposed to solvents, was found to be increased by more than 40%, with no significant heterogeneity across these 7 case-control studies [RE model, OR = 1.43(1.10–1.86),  $I^2 = 16\%$ . (Fig. 3)] or publication bias and small study effects (data not shown). Studies identified in the update specified the nature of the solvents involved. In a study of 101 cases and 110 controls in which measurements of persistent environmental pollutants in whole blood or plasma were made, positive association between ALS and detectable polychlorinated biphenyls 175 (OR = 1.81: 95% CI. 1.20-2.72; p=0.005) and detectable polychlorinated biphenyls 202

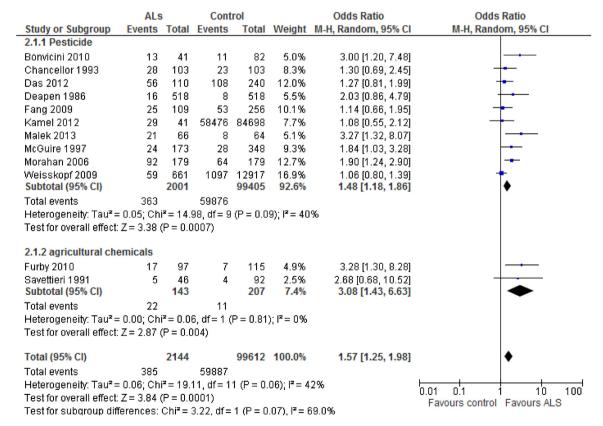


Fig. 2. Previous exposure to pesticides/agricultural chemicals is associated with ALS. The relationship between previous exposure to pesticides (10 articles) or agricultural chemicals (two articles) and ALS was estimated using ReVman 5.1 with random effects model. There was no evidence of significant publication bias (data not shown).

Table 1
Summary of studies related to exposed to solvents and ALS onset.

Author, Year,	Study type	Participants	Characteristics	Solvent
McGuire, 1997, USA	Case control	1.4-year period 1990–1994 2. ALS cases aged 18 or elder years who were newly diagnosed with ALS in western Washington State were identified through a surveillance system, but the cases who lacked a telephone or did not speak English were excluded. 180 cases were eligible, 174 cases agreed to participate. 3.Two controls matched to each case according to sex and age (±5 years) were identified from the study counties using random telephone dialling, or for controls over 65 years of age, Medicare eligibility lists for the target counties was used. 4. Overall response rate is 75%.	1. Cases = 174, controls = 348 2. Men: cases = 95, controls = 190, account for 54.6% in both group 3. White: case = 164, control = 329, account for 94.3, 94.5% respectively 4. Age range not difference, but did not provided average age for both groups 5. Marital status was not different 6. Education level might be different significantly (low/high): cases = 83/91 cases, controls = 124/224 persons.	Solvents: Since there was no agreement for a significant association between self-reported and panel assessed, the results not extracted. But cleaning solvent/degreasers was associated with ALS with agreement: Panel assessment(unexposed/exposed): cases = 131/43, controls = 271/77, OR = 1.9(1.1–3.3)
Chancellor AM, 1993, England	Case control	May 1990 and 31 October 1991, 39 had died before approach was possible.  2. Controls were selected by sex, age, living placed matched	The age and sex of cases and controls were identical, a result of the matching process. There were 61 men (mean age 63, range 35–85 years) and 42 women (mean age 67 range 28–86 years).	Solvents: 1.case/control 5 case only 20 control only 6 neither 72 OR = 3.3(1.3–10). 2. cases (exposed/unexposed) = 25/78, controls = 11/92
Fang F, 2009, USA	Case control	Jacks Match Study participants (cases and controls) were recruited between 1993 and 1996 2. Sequential ALS cases were recruited from two major referral centers in New England. 3. Cases were required to live in New England for at least 50% of the year, to be mentally competent, and to speak English. 71% of eligible cases participated in the study (n = 111). About 85% of the cases were enrolled within 1 year after diagnosis and the remainder within 2 years. 4. Population controls were identified through random telephone screening, no neurological disease, sex, age (30–55, 56–65, and 66–80 years) and region matched. 354 eligible controls were contacted, 270 (76%) were enrolled, 256 completed the entire questionnaire.	1. The median age at diagnosis for cases was 60 years (range, 30–79 years) and the median age at 2 years before interview for controls was 59 years (range, 29–78 years). Diagnosis was made 14.3 months after the onset of symptoms on average.  2. Male: Cases, 66 (60.6); controls, 156 (61.7). Female: Cases, 43 (39.4); controls, 97 (38.3)	Solvents (e.g., toluene or xylene) 1. cases(exposed/unexposed) = 13/96 2. control = 26/237; 3. cases = 109, control = 253.
Qureshi MM, 2006,USA	Case control	1. Between April 1998 and August 2002, recruited 95 subjects with ALS and 106 healthy control subjects in this study. 2. Cases were ide3ntified from clinic. 3. Controls were non-blood relative (spouse), friends, unrelated subjects(age matched).At recruitment controls were matched to the ALS subjects by gender and age.	Cases Controls Gender – Males 60 (63.2%) 58 (54.7%) Race – Caucasian 91 (95.8%) 102 (96.2%) Mean age at enrol. $54.4 \pm 13.1$ (SD) $52.5 \pm 14.9$ (SD) Weight (Kg) – F: $66 \pm 19$ (SD) $65 \pm 10$ (SD) Weight (Kg) – M: $82 \pm 13$ (SD) $86 \pm 17$ (SD)	Solvents (case-control): Of the 95 subjects in the ALS group, the toxin exposure most commonly reported was exposure to pesticides ( $n = 18$ ), followed by lead ( $n = 15$ ), industrial solvents ( $n = 5$ ), mercury ( $n = 5$ ) and miscellaneous toxins.
Savettieri G, 1991, Italy	Case control	Not described for recruitment, no response rate mentioned study period was not included. Controlled were age, sex, social economic status, living place matched	No further information: 46 cases 92 age, sex, residence place, and social economic status matched	Solvents, Exposed/non-exposed: cases = 8/38, controls = 8/84, OR = 2.14(0.6-7.2)
Morahan, 2006, Australia	Case control	1. 179 SALS cases. Did not mention when and how they were selected. Cases have donated	SD = 10 2. Controls: 125 males, 54 females, age, M = 61,	Solvent:  1. OR = 1.76(1.19–2.59) Logistic analysis.  2. Univariate analysis: Solvent/chemical (exposed/non-exposed): cases, 92/86; control, 64/115, OR = 1.92(1.26–2.93); men, cases, 74/51; controls, 55/70, OR = 1.85(1.12–3.04); Women, cases,18/35 controls, 9/45; 2.57(1.05–6.31)
Gunnarsson, 1992, Sweden	Case control	3. Self-reported questionnaires 1. Cases in the age range 45–79 and to a random sample of 500 population controls in the same age range. The questionnaires were answered by 92 cases and 372 controls, a response rate of 85% and 75% respectively. 2. The study population was limited to those aged 45–79 and consisted of 1–2 million inhabitants. 500 controls represented this population. 3. Total of 112 patients diagnosed with MND in the age range 45–79 were identified in the study area.	1. Completed questionnaires were received from 92 cases of MND (58 men and 34 women) and 372 controls (189 men and 183 women), corresponding to a response rate of 85% among the cases (men 83% and women 89%) and 75% among the controls (men 74% and woman 75%).  2. Among the controls the proportion of respondents was proportionately low among those working in agriculture or forestry (63%) but among the cases, this proportion was equal	controls = 65/178, OR = 2.5(1.0-6.5). Trauma(exposed/unexposed): cases = 6/7, controls = 50/79, OR = 1.8(0.5-5.4) 2.Relations (men and women together)

Table 1 (Continued)

Author, Year, country	Study type	Participants	Characteristics	Solvent		
		4. The subjects were asked about their eating habits during the 1960s, 1970s, and 1980s.	to other occupations. The response rate between the different counties ranged from 61 to 83% for the controls and from 74 to 100% for the cases.	(fracture or surgery) less than 10 years before onset of ALS or PBP. No Trauma (exposed/unexposed): cases = 12/24, controls = 65/178, OR = 1.4(0.6-3.0) Trauma(exposed/unexposed): cases = 5/15, CONTROLS = 50/79, OR = 0.7(0.2-2.1).		
Furby, J France	Case control	1. This is a rural area with a rather stable population of about three million inhabitants. All subjects had been living in Brittany for more than 1 year. 2. 2006–2008. 3. The 108 patients, 122 controls were enrolled consecutively from the orthopaedic service of Saint-Brieuc Hospital where they had been hospitalized for minor traumas. Controls were age and sex-matched to the patients. Within the control group, all subjects with chronic neurological disease or alcohol consumption were excluded.	1. Age at interview: Patients, M = 68 SD = 18.0, range[34–86]; Controls, M = 65; SD = 18.0 range[31–84] 2. Sex men/women: Patients, 59/49; Controls, 68/54.	Organic agents or Organic molecules yes/no (reference): Cases = 19/89, controls = 23/99, OR = 0.957 [0.476–1.927], p = 0.903.		
Weisskopf, J 2009, USA	Cohort Study	1. Followed 414 493 male and 572 736 female CPS-II cohort participants who were alive as of 1 January 1989 (earlier ALS deaths were not coded separately), reported no major illness at baseline (1982) and were no missing data on age or sex.  2. ALS deaths were defined as an underlying or contributing cause of death on death certificates. ICD-9 (1989–1998) code 335.2 or ICD-10 (1999–2004) code G12.2 (motor neuron disease) was used.  3. Participants contributed follow-up time from 1 January 1989 to the date of death, or 31 December 2004 (the most recent linkage with NDI), whichever came first.	these analyses were: pesticides/herbicides, 18; asbestos, 10; chemicals/acids/solvents, 36; coal or stone dusts, 8; coal tar/pitch/asphalt, 1; diesel engine exhaust, 9; dyes, 13; formaldehyde, 22;	include organic solvents, a typical misclassification, cohort) Chemicals/acids/solvents No, cases (1014), P.Y = 12,070, Ref, adj.RR1 Ref, adj.RR2 Yes, cases(142), P.Y = 1508, RR = 1.05(0.86–		
Park RM, 2005, USA	Cohort Study	1. Death certificate information for all deaths from 22 participating states in the years 1992–1998 was obtained using the National Occupational Mortality Surveillance System 2. Count Underlying and contributing causes of death 3. Controls were all decedents with no mention of neurologic disease, degenerative or otherwise, and excluding: accidental causes, malignant neoplasms of the brain (ICD 191), other senile and presenile organic psychotic conditions (ICD 290), diseases of the nervous system and sense organs (ICD 320–389), and finally, neoplasms of the lymphatic and hematopoietic tissues (ICD 200–208), due to suspect associations with solvents or electromagnetic fields (EMFs).	1. Motor neuron disease: white men, 3851; white women, 2152; non-white men, 203; non-white women, 141; total, 6,347.2. Total deaths: white men, 1,479,921; white women, 803,110; non-white men, 203,862; Non-white women, 127,453; total, 2,614,346.	Solvents – organic agents(cohort). 1. Solvents: Total deaths 972,505; MND death, 1994, RR = 1.16(1.01–1.34) 2. Benzene: Total deaths, 624,524; 1356, RR = 1.14(0.97–1.33)		
Buckley J, 1983, England	Other	Not mention	Cases = 356+346+335, from mortality data	Exposed to solvent, chemicals, via occupation (cross-section) Leather workers, Clothing workers Construction workers Furnace forge, foundry, rolling mill workers, Textile workers, Painters and decorators, Food, drink and tobacco workers, Professional technical workers, artists, Transport and communications workers, Engineering and allied trades workers, NEC, Sales workers, Paper and printing workers, Mines and quarrymen		

(OR = 2.11; 95% CI, 1.36–3.27; p = 0.001) were observed (Su et al., 2016). In a study of 51 cases and 51 controls in which residential histories were geocoded and linked with databases on air pollution measurements, residential exposure to aromatic solvents was positively associated with ALS (OR = 4.27, 95% CI: 1.09–16.79); (Malek et al., 2015).

3.1.2.4. Exposure to formaldehyde. In an analysis of a cohort comprising about one million participants, an increased relative risk of ALS mortality among adults who reported that they had ever been occupationally exposed to formaldehyde [RR = 1.34 (95% CI: 0.93–1.92) for full cohort; RR = 2.47 (95% CI: 1.58–3.86) after excluding individuals with missing duration of formaldehyde exposure] has been reported (Weisskopf et al., 2009). A strong

	ALS	;	Contr	ol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Chancellor 1993	25	103	11	103	10.3%	2.68 [1.24, 5.79]	
Fang 2009	13	109	26	253	12.0%	1.18 [0.58, 2.40]	<del></del>
Furby 2010	19	108	23	122	13.1%	0.92 [0.47, 1.80]	+
Gunnarsson 1992	12	36	65	243	10.9%	1.37 [0.65, 2.90]	<del> -</del>
McGuire 1997	43	174	77	348	26.7%	1.16 [0.75, 1.77]	<del>*</del>
Morahan 2006	74	125	55	125	21.1%	1.85 [1.12, 3.05]	
Savettieri 1991	8	46	8	92	5.9%	2.21 [0.77, 6.33]	+-
Total (95% CI)		701		1286	100.0%	1.43 [1.10, 1.86]	<b>*</b>
Total events	194		265				
Heterogeneity: Tau² =	0.02; Chi	$i^2 = 7.13$	2, df = 6 (	P = 0.3	1); $I^2 = 16$	%	0.01 0.1 1 10 100
Test for overall effect:	Z = 2.68 (	P = 0.0	107)				Favours control Favours ALS

**Fig. 3.** Previous exposure to solvents is associated with ALS. The relationship between previous exposure to solvent (7 articles) and ALS was estimated using ReVman 5.1 with random effects model. There was no evidence of significant publication bias (data not shown).

dose–response relationship between years of formaldehyde exposure and ALS mortality was identified (trend p=0.0004). The RRs for those reporting <4 years, 4–10 years, and >10 years of exposure formaldehyde to the unexposed were 1.5, 2.1 and 4.1, respectively (Weisskopf et al., 2009). This same research team reported that in a cohort of almost 800,000 men, those who had had jobs with a high probability of exposure to formaldehyde had almost 3-fold greater ALS mortality than in those with no such exposure (Roberts et al., 2015). However, in a cohort study of over 11,000 garment workers exposed to formaldehyde with over 414,000 person-year follow-up, in which there were only eight deaths due to ALS. The association of formaldehyde exposure with ALS mortality was not significant. The major limitation of this study is the sample size (Pinkerton et al., 2013).

3.1.2.5. Previous physical trauma/injury including head trauma/injury. Meta-analysis of 15 case-control studies (Supplementary material VIII) showed a positive association between ALS and

previous trauma, with moderate heterogeneity [RE model, OR = 1.73 (1.43–2.09,  $I^2$  = 53%.] (Fig. 4). Old trauma, defined as trauma that happened at least 5 years prior to ALS diagnosis, was also found to be associated with ALS based on the 4 studies included [RE model, OR = 1.40(1.06–1.86),  $I^2$  = 33%)]. Further subgroup analyses revealed that both previous head trauma [RE model, OR = 1.27 (1.02–1.57),  $I^2$  = 9%] and bone fractures (data not shown) were associated with increased risk of ALS.

These conclusions were further supported by recent results showing a strong association between ALS and physical trauma in a large population based case-control study (Seals et al., 2016) or cerebrovascular injury in a new cohort study (Turner et al., 2015).

3.1.2.6. Lower education. The pooled OR for ALS from three studies (Supplementary material VIII) in individuals with lower education compared to higher education, was 2.04 (95%CI: 1.58–2.62), with no apparent heterogeneity among these three studies (RE model,  $I^2 = 0\%$ ).

	Trauma	ALS	Trauma Co	ontrol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Beghi 2010	29	61	59	112	5.2%	0.81 [0.44, 1.52]	-
Binazzi 2009	42	77	98	185	6.2%	1.07 [0.62, 1.82]	+
Bracco 1979	21	67	10	67	3.5%	2.60 [1.12, 6.07]	<del></del>
Chen 2007	51	105	115	255	7.2%	1.15 [0.73, 1.81]	<del> -</del>
Chio 1991	113	412	72	440	8.9%	1.93 [1.39, 2.69]	-
Deapen 1986	60	518	38	518	7.6%	1.65 [1.08, 2.53]	-
Gallagher 1987	62	135	21	85	5.5%	2.59 [1.42, 4.71]	<del></del>
Gawel1983	24	73	10	61	3.6%	2.50 [1.08, 5.76]	<del></del>
Gresham 1986	38	66	42	66	4.6%	0.78 [0.39, 1.56]	<del></del>
Kondo 1981b	48	158	28	158	6.2%	2.03 [1.19, 3.44]	<del></del>
Kondo 1981F	52	254	28	458	6.7%	3.95 [2.42, 6.45]	_ <del>-</del>
Kondo 1981m	164	458	50	204	8.4%	1.72 [1.18, 2.49]	-
Murros 1983	20	43	55	172	4.7%	1.85 [0.94, 3.65]	<del>  -</del>
Pierce-Ruhland 1981	37	80	28	86	5.2%	1.78 [0.95, 3.35]	<del>  •</del>
Pupillio 2012	225	377	370	754	10.2%	1.54 [1.20, 1.97]	+
Savettieri 1991	15	46	17	92	3.8%	2.13 [0.95, 4.80]	<del>  • </del>
Strickland 1996	10	25	9	50	2.4%	3.04 [1.03, 8.92]	
Total (95% CI)		2955		3763	100.0%	1.73 [1.43, 2.09]	
Total events	1011		1050				
Heterogeneity: Tau <sup>2</sup> = 0		34.21.		0.005): I	²= 53%		
Test for overall effect: Z							0.01 0.1 1 10 100
	2 · · ·	2.220	,				Favours Control Favours ALS

**Fig. 4.** Previous trauma is associated with increased ALS risk. Articles related to trauma including any injury causing medical attention and ALS were searched from multiple literature sources. Data related to injury among ALS and controls from 17 studies in 16 articles was extracted, the synthesized OR was computed based on the odds of injury among ALS cases and controls based on meta-analysis using the random effects model in ReVman 5.1. Similar OR estimates were also obtained with the fixed effects model (data not shown).

Table 2
Summary of studies related to the association of ALS onset with other comorbidities.

Author, Year,	Study type	Participants	Characteristics	Pre-morbidity
Country	.Jr -		,	
Okamoto K, 2009, Japan	Case control	1. Of the 274 ALS patients, 214 (75.3%) eligible cases were enrolled, 183 of whom completed the entire questionnaire. 2.Two community controls that matched to each case for age (G2 years) and gender. 3.Among a total of 732 eligible controls contacted, 550 (75.2%) were enrolled in this study and 430 completed the entire	1. from 2000 to 2005   2. Cases (n 153)   Controls (n 306),60.3% men for cases and controls   3. Mean age (G SD) $63.7 \pm 9.2$ for cases $63.4 \pm 10.6$ controls	BMI and Vigorous activities 1.BMI: $22.2 \pm 0.2$ for cases, $23.3 \pm 0.3$ for control P < 0.05 2. Vigorous activity: $11.6\%$ for cases $6.2\%$ for control OR = $2.0$ ( $1.0-4.0$ ).
Sutedja, J 2011, Hollands	Case control	questionnaire.  1. Between 1 July 2004 and 1 July 2009, patients diagnosed as having sporadic ALS at the University Medical Centre Utrecht.  2. Controls were selected by patients (not spouse, relative, within 5 y, same sex), and also randomly selected by GP using same criteria.  3. 334 patients and 538 controls were included.  4. In the blood sample study, 303 patients and 2100 controls (from other sources) were included.	case, 145(43), for	BMI 1. Compared with controls, fewer patients used cholesterol-lowering agents (OR = 0.6-0.9, p = 0.008) or were overweight (OR = 0.7, 95% CI 0.5–1.0, p = 0.02); 2. Moreover, patients had a lower BMI (ALS BMI = $25\pm3.5$ , control = $26\pm3.6$ , OR = 0.9, 95% CI 0.9–1.0, p = 0.001). 3. TC and LDL were significantly lower, and HDL was significantly higher in ALS patients. 4. The LDL/HDL ratio was significantly lower in patients with ALS (in women, OR = 0.4, 95% CI 0.3–0.6, p < 0.001; in men, OR = 0.5, 95% CI 0.4–0.6, p < 0.001).
Beghi E, Italy, UK, and Ireland	Case control	1. P: have to be definite, probable, or possible ALS	1. Cases/controls = 61/112. P, 27 F, 34 M; C, 46 F, 66 M, controls were not exactly matched. 2. Age, marital, education, no difference between P and C.	BMI BMI. cases: 25.3, SD 4.5; controls: 26.1, SD 4.1 Details were not provided in this paper
Okamoto K, 2009, Japan	Case control	1. Study period: 2000 and 2004 2. Recruited 153 ALS patients aged 18–81 years with disease duration of 3 years.	1. Sex ratio = $60.3/$ 39.7 in both groups 2. Average age: Cases = $63.7 \pm 9.2$ ; control = $63.4 \pm 10.6$	BMI: cases were slimmer than controls: cases = 22.2 $\pm$ 0.3; controls = 23.3 $\pm$ 0.3; P = 0.04

Table 2 (Continued)

Table 2 (Contin	iued)			
Author, Year, Country	Study type	Participants	Characteristics	Pre-morbidity
		3. 306 gender- and age-matched controls were randomly selected from the general population during the study period. 4. Controls: 418 out of 505 completed the entire questionnaires. For cases, should be 100%		
Scarmeas N, 2002, USA		The 431 subjects were consecutive patients seen between 1992 and 2000 in the practice of L.P.R	279 patients with motor neuron diseases and 152 with other neurologic diseases	BMI-slimness subjects with motor neuron diseases were more likely than controls to report they had always been slim or they had been varsity athletes. For slimness(no/yes): Cases = $79/160$ ; control = $61/56$ ; OR = $2.21$ ( $1.40-3.47$ )
Freeman, 2005, USA	Cohort Study	Follow up at least one year	1.9 million population whites only	Comorbidity-cancer 1. There is no significant increase for the risk of ALS among all cancer patients. SMR = 1.0 2. The risk of ALS in melanoma patients is significantly higher than expected SMR = 1.6(1.1–2.6), 32 ALS cases were found in melanoma survivors. The excess incidence of ALS was also observed in tongue tumor survivors SMR = 2.7(1.1–5.7), but only 7 ALS cases were observed.
Turner MT, J 2012, UK		1. Experimental group: with CVD. Oxford Record Linkage Study (ORLS),10 comprising admissions in the former Oxford National Health Service (NHS) region from 1963 to 1998 (ORLS1) and from 1999 to 2008 (ORLS2). 2. Control, no CVD. A cohort without a record of CHD was built consisting of all those admitted to hospital with a wide range of non-cardiovascular conditions comprising mainly minor medical and surgical conditions. This cohort was used as a 'reference cohort', presumed to approximate the general population in its risk of major diseases.	1. 1351 ALS cases in No CHD population from England dataset 2. 1313 ALS cases in CHD population from England dataset 1. 282 ALS cases in no CVD population from ORLS1 and ORLS2 2. 93 ALS cases in CHD population from ORLS1 and ORLS2	1. RR for ALS = 1.14(1.05–1.22) in non-cardiovascular diseases versus vascular diseases, similar results obtained from additional two smaller study populations 2. RR for PD or MS is not different between two groups.
Huisman MHB, 2011, Netherland	Case control Is	1. All newly diagnosed patients and all patients diagnosed before January 2006 and still alive on January 1, 2006, were selected. 2. Controls should be of the same sex and age, plus or minus 5 years. 3. Patients and controls were	1. January 1, 2006, and May 31, 2000. 2. 41 patients (6.4%) had at least one family member with ALS and were, therefore, classified as having FALS. 3. The remainder (594 patients) were classified as having SALS. Relatives of patients have an elevated risk of ALS compared to	decreased, compared to that in the relatives of controls.

Table 2 (Continued)

Author, Year,	Study type	Participants	Characteristics	Pre-morbidi	ity						
Country		asked to fill in structured questionnaires.	controls (any relative 2.42; 95% CI 1.65–3.57).			N	Relatives	Affected	Ratio	OR 95%Ci	
		4. No difference between controls	4. Definite + probable	MI	SALS	594	7679	553	0.072	0.96(0.79-0.94)	
		and patients with regards to the sex ratio, age.	(25), 7 cases are	possible cases: 159		CONT	1566	20925	1749	0.084	
			5. Bulbar onset: 198 (31%) 6. Males: 388(61%) in cases, 935(58%)	CONT ROLS	FALS	41	606	31	0.051	0.61(0.43-0.86)	
			in controls 5. Onset ages: $60.5 \pm 11.4$ ; diagnostic age: $61.8 \pm 11.4$		CONT ROLS	1,616	21584	1805	0.084		
Mitchell JD, 1995, UK	Case control	1. 128 cases were seen in a hospital from a target population of 1.8 million from 1988–1993. 2. 2 controls for each case matched by sex, age with 2 years, and residence. One control with neuron disease, one control was a healthy individual 3. The questionnaires were filled by a nurse	1. Men/women ratio = 1.51 2. Mean age at onset was 61 for men, 64 for women years old respectively.	Comorbidity	y-thyroid d	isease Have	e ever had thy	roid disease: F	RR = 3.02( 1.05	-8.66)	
Pierce- Ruhland, 1981, USA	Case control	1. 88 living patients were identified from hospital records. 80 patients participated the study. 2. Patients friends, same sex, age matched with 5 years were suggested by the patients. 80 controls were contacted, 78 controls	27 women 8 controls: 52 men, 26 women. 8.1 Mean age at interview for both groups were the		-		leurological dis ers in relatives		issociated wil	th ALS. 32 cases, but only	
Sutedja.J 2011, Hollands	Case control	participated.  1. Between 1 July 2004 and 1 July 2009, patients diagnosed as having sporadic ALS at the University Medical Centre Utrecht.  2. Controls were selected by patients (not spouse, relative, within 5 y, same sex), and also randomly selected by GP using same criteria.  3. 334 patients and 538 controls were included.	1. Median age, range: for cases, 60 (24–82); for controls, 59(29–89). 9 female, N(%): For case, 145(43), for controls, 246(46). 9.1 Bulbar onset: 86 (27%) 4. The response rate of the participants in the questionnaire study was 80%	were overweight (OR = 0.7, 95% CI 0.5 – 1.0, p = 0.02);  2. Moreover, patients had a lower BMI (OR = 0.9, 95% CI 0.9 – 1.0, p = 0.001).  3. TC and LDL were significantly lower, and HDL was significantly higher in ALS patients. 4. The LDL/HDL ratio was significantly lower in patients with ALS (in women, OR = 0.4, 95% CI 0.3 – 0.6, p < 0.001; in men, OR = 0.5, 95% CI 0.4 – 0.6, p < 0.001).							

Table 2 (Continued)

Author, Year, Country	Study type	Participants	Characteristics	Pre-morbidity
Mattsson P, 2012, Sweden	Case control	4. In the blood sample study, 303 patients and 2100 controls (from other sources) were included. Linked data: Population Registry, Military Service Conscription Register, Swedish Cause of Death register. Men were born in 1951–1965, and lived in Sweden in	1. Population size: 809 789 total population (122 ALS cases), 686815 has been tested the fitness (85 ALS cases) 2. Nested control were selected	physical fitness(which is determined by genetic factors) ALS is statistically associated with body weight adjusted physical fitness(p=0.01, OR=1.98, 1.32-2.97), but not physical fitness per se (p=0.09, OR=1.01), not body weight(p=0.06, OR=0.98), height(p=0.18, OR=0.98), muscle strength (handgrip, P=0.19, OR=1.0), resting heart rate(p=0.09, OR=1.02).Body weight adjusted knee extension is also associated with increased ALS (p=0.05, OR=1.15(1.00-1.33))
Sørensen, 2010, Finland	Case control	the years of 17th birth day.  1. Performed this population-based case-control study in the northern part of Denmark, which has approximately 1.8 million inhabitants. 2.selected 10 control subjects for each case matched on sex,	Cases Controls Age in years, mean (SD)* 66.2 (12.0) 66.2 (12.0) Male sex, n (%)* 302 (54.3) 3020 (54.3)	Cases Controls Use of statins, n (%) Never 477 (85.6) 4754 (85.5) 1 (reference) Ever 79 (14.2) 806 (14.5) 0.96 (0.73, 1.28) Duration of statin use <3 y 50 (9.0) 455 (8.1) 1.08 (0.78, 1.51) >= 3 y 29 (5.2) 351 (6.3) 0.79 (0.52, 1.20)
Sutedja,J 2011, Hollands	Case control	birth year, and calendar time, from source population  1. Between 1 July 2004 and 1 July 2009, patients diagnosed as having sporadic ALS at the University Medical Centre Utrecht.  2. Controls were selected by patients (not spouse, relative, within 5 y, same sex), and also randomly selected by GP using same criteria.  3. 334 patients and 538 controls were included.  4. In the blood sample study, 303 patients and 2100 controls (from other sources) were included.	1. Median age, range: for cases, 60 (24–82); for controls, 59(29–89). 10 Female, N(%): For case, 145(43), for controls, 246(46). 10.1 Bulbar onset: 86(27%) 4. The response rate of the participants in the questionnaire study was 80%.	Statin-Cholesterol drug 1. Compared with controls, fewer patients used cholesterol-lowering agents (OR = 0.6-0.9, p = 0.008) or were overweight (OR = 0.7, 95% CI 0.5-1.0, p = 0.02); 2. Moreover, patients had a lower BMI (OR = 0.9, 95% CI 0.9-1.0, p = 0.001). 3. TC and LDL were significantly lower, and HDL was significantly higher in ALS patients. 4. The LDL/HDL ratio was significantly lower in patients with ALS (in women, OR = 0.4, 95% CI 0.3-0.6, p < 0.001; in men, OR = 0.5, 95% CI 0.4-0.6, p < 0.001).

3.1.2.7. Strenuous occupation. A meta-analysis of seven case-control studies (Supplementary material VIII) (Fang et al., 2009; Beghi et al., 2010; Binazzi et al., 2009; Gallagher and Sanders, 1987; Kihira et al., 2007; Provinciali and Giovagnoli, 1990; Strickland et al., 1996) showed a statistically significant association between strenuous work and ALS (RE model OR = 2.70, 95% CI: 1.97–3.69), with no significant heterogeneity across the studies.

3.1.2.8. Participating in professional sports, but not recreational sports is associated with ALS. Six case-control studies were included in the meta-analysis of the association between physical activity or participating in sports and ALS [Supplementary material VIII (Okamoto et al., 2009b; Beghi et al., 2010; Longstreth et al., 1998; Vanacore et al., 2010; Veldink et al., 2005b; Pupillo et al., 2012)]. The results showed that participating in recreational sports or

	Eshock	ALS	Eshock Co	ontrol		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Cruz 1999	4	174	4	348	16.0%	2.02 [0.50, 8.19]	+-
Deapen 1986	14	518	5	518	29.7%	2.85 [1.02, 7.97]	<del></del>
Gallagher 1987	8	139	0	85	3.8%	11.05 [0.63, 194.00]	+
Gawel 1983	13	61	5	61	25.9%	3.03 [1.01, 9.12]	-
Savettieri 1991	4	44	2	92	10.4%	4.50 [0.79, 25.58]	<del>  •</del>
Strickland 1996	6	25	3	50	14.2%	4.95 [1.12, 21.84]	-
Total (95% CI)		961		1154	100.0%	3.27 [1.87, 5.73]	•
Total events	49		19				
Heterogeneity: Tau²=	0.00; Chř	= 1.69	, df = 5 (P =	0.89); l²	= 0%		0.001 0.1 1 10 1000
Test for overall effect:	Z = 4.15 (	P < 0.00	001)				Favours control Favours ALS

**Fig. 5.** Previous electric shock is associated with ALS. The association between alcohol drinking (10 articles) and ALS was estimated using the random effects model in ReVman 5.1, providing evidence of a significant increase in risk overall. There was no evidence of significant publication bias (data not shown).

other physical activities was not associated with an increased risk of ALS [RE model, OR = 0.90, 95%CI: 0.78-1.03;  $I^2 = 24\%$ ]. In contrast, ever participating in an organized athletic club in high school or college, or in professional sports, was likely to be associated with ALS (Beghi et al., 2010; Longstreth et al., 1998; Veldink et al., 2005b; Pupillo et al., 2012) [RE model, OR = 1.35, 95% CI: 1.11-1.65.  $I^2 = 3\%$ ].

3.1.2.9. Lower body mass index (BMI). Cohort studies revealed that BMI at baseline was inversely associated with ALS risk (O'Reilly et al., 2013). For each 5-unit increase in BMI within the overweight and obese body weight range, the ALS risk was 21% lower (95% CI: 14-27%). This association persisted among never smokers. Excluding the first seven years of follow-up, the associations remained unchanged, suggesting that weight loss from undiagnosed ALS does not fully explain the findings (O'Reilly et al., 2013). Our meta-analyses from case-control studies supported this inverse association [Table 2, RE model, OR = -0.24(-0.34--0.14),  $I^2 = 0\%$ ].

3.1.2.10. Previous electric shock. The meta-analysis of all 6 case-control studies related to electric shock (Supplementary material IX) showed that previous experience of electric shock was strongly associated with increased risk of ALS [RE model, OR = 3.27 (1.87–5.73),  $I^2 = 0\%$ .] (Fig. 5).

3.1.2.11. Cigarette smoking. The results of a meta-analysis of 20 case-control studies (Supplementary material X) did not support a statistically significant association between smoking and increased risk of ALS [RE: OR=0.97 (0.88–1.08), FE: OR=0.94 (0.87–1.01)] for ever smokers versus never smokers. However, a mild excess relative risk (RR) of developing ALS in female ever smokers compared to non-smokers was observed, based on three cohort studies, with no significant heterogeneity [RE: RR=1.34 (1.17–1.55);  $I^2=12\%$ .]. Other subgroup meta-analyses, such as former versus current smokers among male ALS cases, did not uncover any significant associations between smoking and ALS (data not shown).

However, most studies from Europe and the United States showed that smoking was a risk factor for ALS (Okamoto et al., 2009b; Wang et al., 2008, 2011a,b; Gallo et al., 2009; Kamel et al., 1999; Fang et al., 2006; Berg et al., 2011; Alonso et al., 2010a,b; Schmidt et al., 2010; Sutedja et al., 2007; Weisskopf et al., 2004). A statistically significant association between ever smoking (compared with never smoking) and the risk of ALS was identified in females (RR = 1.7; 95% CI: 1.3–2.1), but not in males (RR = 0.9; 95% CI: 0.7–1.0) using meta-analysis (Alonso et al., 2010a,b). It was estimated that the overall risk of ALS in current smokers (compared with never smokers) was about 1.3 times that of non-smokers (RR = 1.28; 95%CI: 1.0–1.7). This association between

smoking and ALS supports the hypothesis that smoking might be a weak risk factor for ALS (Fang and Ye, 2010; Armon, 2009). However, the inverse (or null) association between lung cancer (Gibson et al., 2016; Catala-Lopez et al., 2014; Freedman et al., 2014), which is strongly associated with smoking, and ALS seems to further support a conclusion of weak association between ALS and smoking, even if there is an association.

3.1.2.12. Military service. As the term 'military service' was excluded when observational studies were screened to identify modifiable risk factors associated with ALS for the general public, we summarize the main observations of the association between ALS and military service here. Excess ALS cases among Gulf veterans, especially aged less than 45 years, were observed in two independent reports, involving comparison with undeployed military personnel (Horner et al., 2003) or the general population (Haley, 2003). The excess cases occurred within the first decade after the Gulf war (Horner et al., 2008), and might be associated with special deployment locations (Miranda et al., 2008). This excess risk of ALS was diminished when using the 13-year follow-up data for the Gulf veterans (Barth et al., 2009).

The risk of ALS due to military service was significantly higher than in controls [RR = 1.53 (1.12–2.09)] (Weisskopf et al., 2005a,b), this observation is supported by following study (Bryan et al., 2016). Although military service itself was not a strong risk factor for ALS [RR = 1.34 (0.87–2.06)] when compared to controls, the RR among military service men for ever participating in war, compared to never participating in war, was significantly increased [RR = 1.36 (1.00–1.71)]. The risk of ALS for war participants increased with the number of wars (RR = 1.57, 1.74, or 1.97) for one, two, and three or more wars, respectively]. Individuals who served during World War II may be at increased risk of ALS (HR: 1.47; 95% CI: 1.13, 1.91) (Weisskopf et al., 2015). The risk of ALS found in a Danish cohort was highest in the decade immediately following the end of military service (Seals et al., 2016).

These observations indicated that the military service associated risk factors, such as chemicals used in battle fields, might be responsible for the increased risk of ALS (Beard and Kamel, 2015).

3.1.2.13. Viral infection. A higher prevalence of spinal fluid enterovirus genome has been detected in ALS patients in some studies (Vandenberghe et al., 2010; Cermelli et al., 2003; Berger et al., 2000; Jubelt, 1992), but not in others (Jubelt and Lipton, 2004; Nix et al., 2004). An association between polio virus infection and ALS has also been proposed (Esik et al., 2004), but has never been confirmed.

3.1.2.14. Cardiovascular factors. Genetic variants of VEGF, the gene encoding a growth factor for vascular endothelial cells, have been

Table 3
Summary of studies related to alcohol intake and ALS onset.

Author, Year, Country	Study type	Participants	Characteristics	Others
Nelson LM, 2000, USA	Case control	1. Control: randomly selected. The combined response rate for eligible controls identified and contacted by using both methods was 79 percent (348/441).  2. Patients had to be residents of one of the three counties, aged 18 years or older, and newly diagnosed with ALS during the 4- year study period.	N/A	Alcohol coffee 1. Alcohol: no association had been identified. cases(yes/no) = 73/88, controls = 129/192.  2.No relation was observed between caffeine intake and the risk of ALS (relative to the lowest quartile of caffeine intake, the odds ratios adjusted for age, gender, respondent type, education, and smoking were 0.9, 1.0, and 0.9 for the second through fourth quartiles of
Okamoto K, 2009, Japan	Case control	1. Of the 274 ALS patients, 214 (75.3%) eligible cases were enrolled, 183 of whom completed the entire questionnaire.  2. Two community controls that matched to each case for age (G2 years) and gender.  3. Among a total of 732 eligible controls contacted, 550 (75.2%) were enrolled in this study and 430 completed the entire questionnaire.	1. From 2000–2005	caffeine intake, respectively). Alcohol drinking, stress, type A behaviour, less intake of green vegetables 1. Current drinking: 69.1% for cases, 68.0% for controls OR = 1.1 (0.7-1.5). Cases (yes/no) = 106/47, control = 208/98. 2. Self-assessed stress: 61.9% for cases, 47.4% for controls OR = 2.9 (1.3-2.7) 3. Type A: 44.2% for cases 19.6% for controls OR = 2.9 (1.9-4.5) 4. Interaction between green/yellow vegetables: 50.8% for cases 70.6% for controls OR = 2.5 (1.7-3.7) 5. Not confounded between type A and vegetable consumptions.
Kamel F, 2002, USA	Case control	Has been described in other papers	No difference between two groups with regard to the age, sex ratio, education level	
Okamoto K, Japan, 2009	Case control	1. Study period: 2000 and 2004 2. Recruited 153 ALS patients aged 18–81 years with disease duration of 3 years. 3. 306 gender- and age-matched controls were randomly selected from the general population during the study period. 4. Controls: 418 out of 505 completed the entire questionnaires. For cases, should be 100%	1. Sex ratio = 60.3/39.7 in both groups 2. Average age: Cases = 63.7 $\pm$ 9.2; control = 63.4 $\pm$ 10.6	Alcohol drinking, 1. Drinking: cases = 35.3%, controls = 31.4%. cases(yes/no) = 108/198, controls = 131/287.
Beghi E, 2010, Italy, UK, and Ireland	Case control	1. P: have to be definite, probable, or possible ALS cases, did mention when the study was done 2. Age ( $\pm 2.5$ y) and sex matched for control for each patient, selected by general practitioners for the patients 3. Report response rate is more than 90%, no further data	1. Cases/controls = 61/112. P, 27 F, 34 M; C, 46 F, 66 M, controls were not exactly matched. 2. Age, marital, education, no difference between P and C.	Alcohol  1. Work intensity (Yes/no, $P < 0.05$ ): Strenuous/ all, $P = 8/53$ , $C = 4/108$ ;  2. Sport intensity (yes/no, $P > 0.05$ ): strenuous/ all others, $p = 8/53$ , $C = 19/93$ 3. Professional sport(yes/no, $p = 0.04$ ): $P = 3/58$ , $C = 0/112$ 4. Alcohol(yes/no): $P = 18/43$ ; $P = 18/43$
Kondo K, 1981, Japan	Case control	1. Cases were identified from multiple sources, including death certificates. 2. Spouses were served as controls 3. Spouses who have lived with patients since their marriages provided information. If patient who had no spouse to provide this information, or unmarried patients were excluded 4. Spouses were interviewed by a visiting nurse for the events from marriage to the disease onset (not diagnosis) Study B, case-control, age, sex, and residence matched, but the patients and controls were interviewed by neurologists in study B.	1. Cases: 458 men, 254 women. Controls, their wives, or husbands. 1965–1966 2. Study B:1973	
Savettieri G, 1991, Italy	Case control	Not described for recruitment no response rate mentioned study period was not included Controlled were age, sex, social economic status, living place matched	No further information 46 cases 92 age, sex, residence place, and social economic status matched	Alcohol(exposed/non-exposed): cases = 18/28, controls = 37/55, OR = 0.76(0.4–1.4)
Fang F, 2009, USA	Case control	1. All study participants (cases and controls) were recruited between 1993 and 1996 2. Sequential ALS cases were recruited from two major referral centers in New England. 3. Cases were required to live in New England for at least 50% of the year, to be mentally competent, and to speak English. 71% of eligible cases participated in the study (n = 111). About 85% of the cases were enrolled within 1 year after diagnosis and the remainder within 2 years.		agents Overall smokers non-smokers 1. Exposed to Adhesive: 36 23 1.5 (0.8–2.7) 22 15 1.2 (0.5–2.6) 14 8 2.4 (0.9–6.9) 2. Dyes or printing inks: 37 21 1.4 (0.8–2.7) 22 13 1.1 (0.5–2.4) 15 8 2.5

Table 3 (Continued)

Author, Year, Country	Study type	Participants	Characteristics	Others
		4. Population controls were identified through random telephone screening, no neurological disease, sex, age (30–55, 56–65, and 66–80 years) and region matched. 354 eligible controls were contacted, 270 (76%) were enrolled, 256 completed the entire questionnaire.		(1.9–20.4) 4. Antifreeze or coolants: 23 17 1.6 (0.8–3.3) 17 10 0.9 (0.3–2.4) 6 7 6.3 (1.6–24.1) 5. Degreasers or cleaning agents: 59 28 1.0 (0.6–1.7) 40 20 0.8 (0.4–1.5) 19 8 1.5 (0.5–4.2) 6. Mineral spirits/white spirits: 20 13 1.7 (0.8–3.7) 13 7 1.0 (0.4–2.8) 7 6 4.1 (1.2–14.4) 7. exposed to formaldehyde: Never, controls 204, cases 89; Ever, controls 49, cases 20, OR=0.8(0.5–1.5)
kamel F, 1999, USA	Case control	1. Cases were recruited from two referral center between 1993–1996. 2. Cases were eligible to participated if they had received their first diagnosis within two years.3. Speak English and able to participate the study, i.e., no mental and physical disability.  4. 223 sequential cases, 154 cases were eligible, and final 109 cases finished the studies. 5. Population Controls were identified through telephone screening. 354 eligible controls were contacted, 270 agreed to participated, and 256 finish the questionnaires	1. Conducted between 1993–1996 2. 109 cases, 256 controls, 3. 95% whites 4. 61% men, 39 women in both cases and controls 5. No average age, but case and control were age matched, the ages range from 35 to 80 in both groups.	Alcohol: no association: cases (yes/no) = 29/71, controls = 42/214.
de jong, Holand, 2012	Case control	1. 494 incident ALS patients diagnosed with ALS on or after January 1, 2006. 2. 255 prevalent ALS patients diagnosed before January 1, 2006, but alive after that date. 3. An overlap of 178 patients between the prevalent and previously studied patient groups. 4. A referral ALS population (n = 366 or 359) diagnosed between January 1, 2001, and December 31, 2005.	and prevalent ALS patients in terms of sex ratio, age at onset, disease duration and education level. 2. Bulbar type in incident (34.8%) versus prevalent (24.8%). 3. Elementary education level in control (5.7%) versus in all ALS patients	Incident case(494) prevalent cases(255) control (1599) never 15.8 20.2 9.7 Former drinkers 5.3 7.4 5.1 Current 78.9 72.4 85.2

associated with ALS (Terry et al., 2004; Barbeito et al., 2010; Lambrechts et al., 2009; Wang et al., 2007). Beneficial vascular factors, such as lower LDL/HDL, have been associated with an increased risk of ALS (Sutedja et al., 2011; Turner et al., 2012; Kim et al., 2011; Paganoni et al., 2011, 2012; Dorst et al., 2011).

3.1.2.15. Beta-N-methylamino-L-alanine (BMAA). ALS clusters in the Gulf desert and in New Hampshire have been attributed to the intake of BMAA (Cox et al., 2009; Caller et al., 2009, 2012). Over the last few years, several ecological studies showed an association between the spatial distribution of BMAA and the incidence of ALS in France (Masseret et al., 2013) and the USA (Field et al., 2013).

3.1.2.16. Alcohol. A meta-analysis of existing observational studies showed that alcohol drinking was not associated with increased ALS risk (data not shown, Table 3). However, a recent study indicated an inverse association (incident patient group: OR = 0.52, 95% CI: 0.40, 0.75) (de Jong et al., 2012), suggesting alcohol may be a protective factor against ALS.

3.1.2.17. Coffee. Based on a case-control study of 377 cases and three control groups, drinking coffee may be a protective factor against ALS (Beghi et al., 2011). The ORs for ALS comparing to neurologic, non-neurologic, and controls recruited through family practices were 0.7 (95% CI: 0.5, 1.0), 0.5 (95% CI: 0.3, 0.7), and 0.4 (95% CI: 0.2, 0.8), respectively (Beghi et al., 2011). However, this observation was not seen in a pooled analysis of data on 1279 cases

**Table 4** Monogenic mutations related to ALS onset.

Gene	Protein name and function	transmission	Gene location	Mean onset age (y)
ALS2	Amyotrophic lateral sclerosis 2	AD	2q33.2	1
C9orf72	Unknown	AD	9p21.2	
DCTN1	Dynactin 1 (p150, glued homolog, Drosophila)	AD	2p13	55
FUS	Fusion (involved in t(12;16) in malignant liposarcoma)	AD	16p11.2	45
OPNT	Optineurin	AD	10p13	51
SETX	Senataxin, ALS 4	AD	9q34.13	18
SOD1	Superoxide dismutase 1, soluble (amyotrophic lateral sclerosis 1 (adult)	AD/AR	21q22.11	47
SPG11	Spastic paraplegia 11	AR	15q14	16
TAF15	TAF15 RNA polymerase II, TATA box binding protein (TBP)-associated factor, 68 kDa		17q11.1-q11.2	50
TARDBP	TAR DNA binding protein	AD	1p36.22	55
UBQLN2	Ubiquilin	X-linked	Xp11.21	41
VAPB	VAMP (vesicle-associated membrane protein)-associated protein B and C	AD	20q13.33	44

occurring in over one million people included in 5 cohort studies (Fondell et al., 2015).

3.1.2.18. Vitamin E, vitamin C and unsaturated fatty acids. Intake of supplemental vitamin E, but not other vitamins, has been associated with reduced risk of developing ALS, which is consistent with ALS being related to reactive oxygen species (Veldink et al., 2007; Wang et al., 2011a,b). A case-control study showed that intake of vitamin E was associated with a reduced ALS risk (OR = 0.4, 95% CI: 0.2-0.7, p = 0.001) (Veldink et al., 2007). A large pooled prospective study showed that an inverse doseresponse between dietary vitamin E intake and ALS risk was evident in women, but not in men (Wang et al., 2011a,b). The most recent cohort study showed that the age-adjusted relative risk for ALS was significantly lower in subjects with higher serum Vitamin E levels (RR = 0.56, 95%CI: 0.32-0.99, p = 0.046) when compared to those with lower levels (Michal Freedman et al., 2013). In addition, intake of vitamin C, carotenoids, and unsaturated fatty acids may also reduce ALS risk, or delay ALS onset (Fitzgerald et al., 2013, 2014).

3.1.2.19. Fruit and vegetables. Two case-control studies from Japan reported an inverse association between intake of fruit/vegetables and the risk of ALS (Okamoto et al., 2009a,b).

3.1.2.20. Chicken soup. One cohort study showed that the risk of ALS, among people who consumed chicken soup versus never having consumed chicken soup, was significantly lower after adjusting for age and smoking [RR=0.94 (0.66–1.34) for one serving, 0.79 (0.55–1.13) for 2–3 servings, and 0.58 (0.36–0.94) for more than 4 servings per day], with a statistically significant dose-dependent reduction trend (trend p=0.0006) (Morozova et al., 2008).

## 3.1.3. Genetic factors

Many genetic risk factors have been associated with ALS. Several internet sites have been established to collect information from genetic studies of ALS (He et al., 2015). The most comprehensive web site is the ALSOD site (Lill et al., 2011; Abel

et al., 2012). It not only collects complete gene mutation information and basic demographic information of patients, but also provides researchers with some simple data analytic results and related lists of literature references. Most of the genetic information presented below was summarized from this website. However, we also ran an article search for some risk factors (the search strategy for genetic factors is outlined in Supplementary material III).

factors a Mendelian 3.1.3.1. Genetic with inheritance pattern. Although ALS is a relatively rare neurodegenerative condition, 658 mutations in 126 genes have been linked to ALS. ALS mutant genes could be found in every chromosome, except Y chromosome. Of these 126 genes, 32 mutated genes (Table 4) may be transmitted in a Mendelian inheritance pattern based on 1094 ALS patients with complete genetic data (http://alsod.iop.kcl.ac. uk/Statistics/report.aspx, access 2016), explaining about 10% of diagnosed ALS cases (Chio et al., 2012). This list of gene mutations is growing (Abel et al., 2012). The two most common monogenic mutations are of SOD1 and C9orf72 among white ALS patients, accounting for 25% and 35% of familial cases, and 2% and 6% of sporadic cases, respectively (Chio et al., 2012; Wijesekera and Leigh, 2009; Pfister et al., 2013; Robberecht, 2000; van Blitterswijk et al., 2012b). C9orf72 mutations have been reported in populations of European origin, but very rarely in Chinese or Japanese populations (Turner et al., 2013; van Blitterswijk et al., 2012b; Chen et al., 2016). The C9orf72 mutation is thought to have arisen in Scandinavia several thousand years ago (Smith et al., 2012; Ratti et al., 2012). The manner by which environmental factors might interact with this gene has not been well investigated (Figueroa-Romero et al., 2012), but has suggested that the expression of these genes might be modulated by environmental risk factors via DNA methylation (such as on FUS gene) (Kaneb et al., 2012). Unlike most inherited diseases, the onset of inherited ALS usually occurs late in life, at around 50 years of age [calculated from ALSOD website (ALSOD, Jan. 2014)], and can occur as late as 85 years of age (Aggarwal and Nicholson, 2005). The inheritance penetration is much lower than 100% for most genes of this type. Therefore, the onset time of inherited ALS might be subject to modification by

**Table 5**Meta-analyses for the associations between gene variants and ALS onset.

Gene	Chromosome site	Name Function	Variants	Studies included	Pooled OR	Conclusion
ANG	14q11.1	Angiogenin Functionally similar to VEGF	Rs11701 (G/T)	3 p, 2 n, 5 no association	1.13(0.93– 1.40) i2=75	Heterogeneity is high across studies, conclusion might be not appropriate
APEX1	14q11.2	APEX nuclease (multifunctional DNA repair enzyme) 1	Rs1130409 (T/G)	1 n, 3 no association	0.78(0.62– 0.97) i2 = 20	Gene variant may reduce the risk for ALS
APOE ATXN2	19q13.2	Apolipoprotein E	Rs429358 (C/T)	1 n, 8 no association	0.94(0.79– 1.12) i2=44	May be associated with age of onset, presentation and survival Strongly associated with it
GRIA2	4q32.1	Glutamate receptor	Rs10025251 (C/T)	6 no association	0.99(0.91– 1.08) i2=0	Seems no association with this variant
GWA_9p21	9p21.2	unknown	Rs2814707 (A/G)	6 p, 1 n, 9 no association	1.24(1.18– 1.30) i2 = 70	Associated with ALS, but heterogeneity is high across studies. Also associated with dementia
HFE	6p21.3	Haemochromatosis oxidative stress	Rs1799945 (G/C)	4 p, 2 no	1.72(1.20– 2.48) i2 = 74	Contrary to prediction, this variant is associated with ALS, although heterogeneity across studies
PON1	7q21.3	Paraoxonase 1 metabolism of xenobiotics		1 p, 7 no association	1.10(1.00– 1.22) i2 = 32	This variant is associated with ALS, but inconsistent with a recent <i>meta</i> -analysis. I favour this association here.
PON2	7q21.3	Paraoxonase 1 metabolism of xenobiotics	Rs7493 (C/ G?)	1 p, 5 no association	1.10(0.86– 1.42) i2 = 34	May not be strongly associated with variant
PON3	7q21.3	Paraoxonase 1 metabolism of xenobiotics	Rs10487132 (G/A)	1 p, 4 no association	1.09(0.97– 1.23) i2 = 44	May not be strongly associated with variant
UNC13A	19p13.11	protein unc-13 homolog A, Neurotransmitter release	Rs12608932 (C/A)	8 p, 9 no association	1.15(1.10– 1.21) i2 = 31	Also a survival modifier. Diekstra FP, Neurobiol Aging. 2012 Mar; 33(3): 630.e3-8
VEGFA	6p12	Vascular endothelial growth factor	Rs699947	1 p, 9 no association	1.06(0.98– 1.16) i2=0	May not be strongly associated with ALS

Table 6
Studies related to PolyQ expansion in ataxin-2 gene.

Author, Year, Country	ALS or MND Diagnosis	Sporadic, Familial sex, Case, Control and PolyQ information	Main results and Conclusion
Included studies Conforti, 2012, Italy	ALS, EI Escorial, exclude SOD1, TDP-43, ANG, FUS, C9ORF72 positive, Control matched with geographic	1. 405 sALS: 2. 13 fALS. 3. 296 control:	1. 57 out of 806 <i>ATXN1</i> alleles in sALS cohort harboured a ≥ 32 polyQ repeat (7.07%), compared to 13 out of 544 NC alleles (2.38%, p=0.0001). OR = 2.396(1.26-4.56) 2. For <i>ATXN2</i> , 22 X ≥28Q in 808 sALS alleles (2.72%) and only 3 (0.5%) of 586 NC alleles (p=0.01). OR=5.832(1.71=9.78).
Corrado, 2011, Italy	ALS, EI Escorial. Screened SOD1, TDP-43, FUS, regionally matched controls	1. 232 ALS (219 sALS, 13 fALS): $1 \times 24Q$ ; $1 \times 27Q$ , $0 \times 28Q$ ; $1 \times 29Q$ ; $0 \times 30Q$ ; $3 \times 31Q$ ; $1 \times 32Q$ ; $2 \times 33Q$ ; $1 \times 37Q$ . 2. 395 controls: $3 \times 24Q$ ; $6 \times 27Q$ , $1 \times 28Q$ ; $4 \times 29Q$ ; $1 \times 30Q$ ; $0 \times 31Q$ ; $0 \times 32Q$ ; $0 \times 33Q$ ; $0 \times 37Q$ .	(p-0.01). (p-3.032(1.71-3.70).  1. Significantly association.  2. No observation for symptoms of ataxia, dementia or other atypical features.
Daoud H, 2011, France/Qubec, Canada	1. ALS, EI Escorial, probable, or definite cases, diagnosed by ALS specialists. FALS were sod1, tardbp, fus, vapb, ang free. Neurological health controls matched by age, ethnicity. Not mentioned sex, and recruitment time.	1. 461 sALS: ≥32, 9; 2. 95 fALS: ≥32, 2; 3. 471 unrelated controls. Range from 19–32. 19– 37 in patients. 24 of 471 healthy controls (5,1%) harboured 1 intermediate ATXN2 allele(range 24–33), whereas 40 of 556 cases (7.2%) had one allele in that range (p = 0.15). 4. Figure may help do scientific guess.	1, ROC showed that 27 cut-off gave best sensitivity/specificity. 19 from controls (4,1%), compared to 35 from cases(6.3%) within this range ( $\geq$ 27), no difference(p=0.09). The p becomes significant when compared within the range of $\geq$ 29. 4 controls (0.8%), and 25 cases (18/461, 4.5%). OR=5.5(1.9–15.9). 7 out of 95 fALS (7.4%), 18 out of 461 sALS cases: OR for FALS was 9.29(2.66–32.4); for sALS is 4.74(1.59–14.13). When $\geq$ 32 remove, the difference is still significant (14 from cases, 3 from controls, OR=4.03(1.15–14.11)) for the range between 29 and 31. 2. Length is not associated with age at onset ( $\geq$ 29) in cases were no $\geq$ 29 cases.
Elden, 2010, USA	ALS, recruited in PA and NJ, no recruitment time, no case control design	1. 915 ALS. $6 \times 24Q$ ; $1 \times 25Q$ ; $0 \times 26Q$ ; $22 \times 27Q$ ; $1 \times 28Q$ ; $2 \times 29Q$ ; $4 \times 30Q$ ; $7 \times 31Q$ ; $8 \times 32Q$ ; $2 \times 33Q$ . 2. 980 Controls: $4 \times 24Q$ ; $4 \times 25Q$ ; $2 \times 26Q$ ; $11 \times 27Q$ ; $0 \times 28Q$ ; $1 \times 29Q$ ; $0 \times 30Q$ ; $2 \times 31Q$ ; $0 \times 32Q$ ; $0 \times 33Q$	<ol> <li>CAG intermediate repeat expansion is associated with ALS.</li> <li>The repeat might be a modulator.</li> </ol>
Gellera, 2012, Italy	ALS, EI Escorial, SOD1 free, for all, FALS also screened ANG, TDP-43, FUS, C9ORF72 (carriers of FALS cases-separated)	1. 658 SALS: $4 \times 24$ -26Q; $10 \times 27Q$ , $4 \times 29Q$ ;	1. The frequency of ATXN2 alleles with 27–30 repeats was similar in SALS and control subjects. 2. Fifteen SALS subjects carried ≥ 31 CAG repeats. This difference was statistically significant (p=0.0014). No alleles with ≥ 34 CAG were found. 3. In FALS, the distribution of ATXN2 alleles was similar to control subjects.
Gispert, 2012, Germany	ALS, European ALS clinics (diagnostic criteria),	1. 559 ALS (included 89 fALS, 1 × 32Q): 1 × 24Q, 1 × 25Q; 0 × 26Q; 15 × 27Q; 0 × 28Q; 1 × 29Q; 3 × 30Q; 0 × 31Q; 3 × 32Q; 0 × 33Q; 0 × 34Q; 1 × 35Q. 2. 1378 controls: 9 × 24Q, 1 × 25Q; 0 × 26Q; 48 × 27Q; 1 × 28Q; 5 × 29Q; 1 × 30Q; 1 × 31Q; 0 × 32Q; 0 × 33Q; 1 × 34Q; 0 × 35Q. 3. 1142 PD: 6 × 24Q, 2 × 25Q; 1 × 26Q; 37 × 27Q; 1 × 28Q; 5 × 29Q; 0 × 30Q; 2 × 31Q; 1 × 32Q; 0 × 33Q; 0 × 34Q; 3 × 37Q; 1 × 39Q; 1 × 40Q.	1. In 559 sporadic ALS patients from Central Europe, the association of ATXN2 expansions (30 ≤ polyQ ≤ 35) with ALS was highly significant. 2. The study of 1490 patients with Parkinson's disease (PD) showed an enrichment of ATXN2 alleles 27/28 in a subgroup with familial cases, but the overall risk of sporadic PD was unchanged. 3. No association was found between polyQ expansions in Ataxin-3 (ATXN3) and ALS risk.
Lahut, 2012, Turkey	ALS, no other information	1. 236 ALS. $4 \times 24Q$ ; $1 \times 25Q$ ; $3 \times 27Q$ ; $1 \times 28Q$ ; $1 \times 29Q$ ; $1 \times 31Q$ ; $3 \times 32Q$ . 2. 420 control. $3 \times 24Q$ ; $0 \times 25Q$ ; $1 \times 27Q$ ; $0 \times 28Q$ ; $0 \times 29Q$ ; $0 \times 31Q$ ; $0 \times 32Q$ .	capalistoris in AdaMir-3 (ATAMS) and ALS 118A.  1.15 ALS patients carrying SOD1, UBQLN2, OPTN, SPG11, or PLEKHG3 were intermediate repeat negative.  2. 4 X > 31Q (1fALS, 3sALS). Could not find any information about total FALS cases.  2. Calculate > 30 in this study. 3 out of 4 patients with 31 and 32Q had a single CAA interruption. Not in other ALS cases.
Lee T, 2011, Multiple European countries	ALS, El Escorial, Controls matched by age and gender, who were either the spouses of ALS patients, healthy donors. Did not mentioned recruitment time	1. 400 fALS: >30, 6; 2. 894 sALS: >30. 7; 3.679 controls. 20 (2.9%) out of 679 controls harboured intermediate polyQ(range 27–30). 45 out of 1294 ALS patients (3.5%, range 27–35). For >30 repeat, no in 679 controls, but found 14 cases among 1294 ALS patients (p=0.0062).	1. No ataxia, dementia was observed in ALS patients. 2. No difference compared with and without repeat for age at onset, disease duration. 3. Intermediate-length ataxin 2 polyQ repeat expansions are associated with increased risk for ALS also in the European cohort. The specific polyQ length cut-off, however, appears to vary between different populations, with longer repeat lengths showing a clear association.
	1. El Escorial. Recruited from 2008–2010		tengens showing a clear association.

Table 6 (Continued)

Author, Year, Country	ALS or MND Diagnosis	Sporadic, Familial sex, Case, Control and PolyQ information	Main results and Conclusion
Ross, USA, Canada, 2011		1. 532 als: $\geq$ 27Q, 33(6.2%); $\geq$ 31Q, 8(1.5%). 2. 4877 control: $\geq$ 27Q, 197(4.0%); $\geq$ 31Q, 9(0.2%). 3. 642 FTD: $\geq$ 27Q, 31(4.8%); $\geq$ 31Q, 9(0.5%). 4. 1530 CE: $\geq$ 27Q, 56(3.7%); $\geq$ 31Q, 3(0.2%). 5. 514 PSP: $\geq$ 27Q, 24(4.7%); $\geq$ 31Q, 4(0.8%). 6.702 PD: $\geq$ 27Q, 28(4.0%); $\geq$ 31Q, 2(0.3%).	1. ALS, for $\geq$ 27, OR = 1.58(1.08–2.31); for $\geq$ 31, OR = 5.57(1.95–15.88). 2. FTD, for $\geq$ 27, OR = 1.20(0.82–1.76); for $\geq$ 31Q, OR = 1.94(0.51–7.37) 3. for AD, for $\geq$ 27, OR = 0.96(0.70–1.33); for $\geq$ 31Q, OR = 2.17(0.40–11.96) 4. For PSP, for $\geq$ 27, OR = 1.20(0.78–1.85); for $\geq$ 31Q, OR = 5.83(1.74–19.52) 5. For PD, for $\geq$ 27, OR = 0.95(0.63–1.43); for $\geq$ 31Q, OR = 0.93(0.19–4.51). Author speculated that long Q repeat in controls (9 controls) might be due to young age (reduced disease penetrance). SCA2 not CAA interruption?
Soraru, 2011, Italy	ALS, EI Escorial, recruited from 01 of 2004–08 of 2010. Did not mention control match	1. 247 ALS. $\geq$ 24Q, 17(6.8%). 3 × 24Q, 1 × 26Q, 6 × 27Q, 2 × 30Q, 1 × 31Q, 4 × 32Q 2. 256 control: $\geq$ 24Q, 6(2.3%). 1 × 24Q, 2 × 27Q, 1 × 28Q, 2 × 31Q.	Intermediate polyQ is more frequent in ALS patients, than in controls (p = 0.026).     No difference was observed for age at onset, bulbar/spinal onset ratio, survival time, etc.
Van Damme, 2011, Belgium/ Netherlands	1. ALS, EI Escorial. 1995–2010, neurological conditions free normal controls	1. 1845 SALS. ≥32, 10(5 × 32, 2 × 33, 1 × 34, 1 × 36, 1 × 39; 0.5%); (31,4; 30, 5; 29,9; 28, 1; Scientific guess from fig. we can further guess 27 repeat) 2. 103 fALS cases from 91 families. but sod1, fus, TARDBP, ANG free. 2/91 (2.2%) long repeat, 1/91, 31 repeat; 1/91, 33 repeat; 3. 2002 controls. Range, 16−31; 22, 90.1%; 23, 6.1%; 27, 1.7%; 31, 0.1% (heterozygous, 0.2%). (31,5; 30, 4; 29,7; 28, 1; Scientific guess from fig.)	1. p = 0.0006 for repeat ≥32 between ALS and control. No difference for < = 31 (22–31, or 27–31 or 29–31).  2. ROC curve show a cutoff ≥29 yield greatest sn, spn.28 out of ALS patients (1.5%) versus 16 out of 2002 controls (p = 0.036, OR = 1.92(1.04–3.64)). Combined with an American study (915 ALS, 980 controls), OR = 2.93(1.73–4.98).  3. No association with survival, age at onset, site of onset.  4. Pedigree (33:33), onset at 71, and his 2 y elder als brother (31:33, onset at 75), normal brother (22:33) were described from consanguineous family. One of their parents was possible affected by ALS. No ataxia or cerebellar degeneration was found.
Studies from Chi Chen, 2011, China	na 1. ALS, EI Escorial, 05-2004-06-2010. Excluded fALS, Community controls matched by age, sex, race from same period.	1. 345 sALS (254 spinal, 91 bulbar). ≥24, 15; <24, 330; ≥27, 12; <27, 333; ≥28, 11; ≥29, 8; ≥31, 4. 2. 350 controls (17–30). ≥24, 8; <24, 342; ≥27, 4; <27, 346; ≥28, 3; ≥29, 2; ≥31, 0. 3. Provide a table for comparing repeat length	between with and without ATXN2, no difference.
Liu, 2013, China	ALS, El Escorial, whole China, did not state the recruitment time, no control match information	and clinic features 1. 1067 als: 4×26Q; 4×27Q; 3×28Q; 6×29Q; 3×30Q; 17 X>30Q (6×31Q; 5×32Q; 3×33Q; 2×34Q; 1×35Q). 2. 506 healthy: 1×26Q; 2×27Q; 4×28Q; 0×29Q; 7×30Q; 0X>30Q. 3>6 fALS, no.	Association of ALS with ataxin-2 intermediate CAG repeats was confirmed.     No clinical manifestation associated with repeat observed.
Excluded article Van Langenhove, 2012, Belgium	No diagnostic, recruitment and control information.	1. 72 ALS cases, including 18 fALS. 27–33Q, 7; 30–33Q, 3; >31Q, 1 (33Q). No information for fALS alone. 2. 22 FTLD-ALS: 27–33Q, 1; >31Q, 0. 3. 270 FTLD. 27–33Q, 8; 4. 810 controls.27-33Q, 25; >31, 0.	and SCA2 was observed.  2. Intermediate repeat of polyQ is associated
Multiple use of s Bonini, 2011, USA	ame sample Same as Elden	Title Model organisms reveal insight into human neurodegenerative disease: ataxin-2 intermediate-length polyglutamine expansions are a risk factor for ALS	Main results 1. For 27–33Q; 1.4% in control; 4.7% in ALS (including sALS and fALS). Significantly associated intermediate polyQ repeats with ALS. 2. Mentioned that the cut-off appeared dependent on the specific population. 3. This study used similar approaches as in Elden's paper
Lee T, 2011, USA	Same as Elden	Evaluating the prevalence of polyglutamine repeat expansions in amyotrophic lateral sclerosis	1. Assessed the polyQ lengths of ataxin 1, ataxin 3, ataxin 6, ataxin 7, TBP, atrophin 1, and huntingtin in several hundred patients with sporadic ALS and healthy controls.  2. Other than ataxin 2, we did not identify a significant association with the other polyQ genes and ALS
Yu Z, 2011, USA	Same as Elden	PolyQ repeat expansions in ATXN2 associated with ALS are CAA interrupted repeats	Expanded repeat alleles of 40 ALS patients and 9 long-repeat length controls were all interrupted, bearing 1–3 CAA codons within the CAG repeat.

environmental factors. Several rarer gene mutations causing the onset of ALS in children have also been reported (ALSOD, Jan. 2014)

3.1.3.2. Genetic polymorphisms. Identifying the associations with polymorphic genetic variants in individuals with ALS has been an important research topic for decades. In addition to the ALSOD website, the ALSgene website mainly focuses on the information of gene polymorphisms identified in ALS cases. This website provides a list (total number is 22 as of May 2016) of top polymorphisms with strong association (P < 0.0001) with ALS concluded by the results of at least one meta-analysis (Table 5). Of these 22 polymorphisms, 14 of them may increase ALS risk with mean OR between 1.12 and 1.56, whereas remaining 8 tested polymorphisms may decrease ALS risk with mean OR between 0.66 and 0.88. Establishing replicated associations between ALS and a gene polymorphism has proved difficult. For example, the association between the polymorphisms in gene PON1 reported in a few early studies has not been confirmed by a large scale metaanalysis based on pooling samples from multiple groups, including the samples from those early studies (Wills et al., 2009); this same conclusion was reached in subsequent studies (Ricci et al., 2011; van Blitterswijk et al., 2012a; Chen et al., 2012; Lee et al., 2015). Similar inconsistent results regarding the association between polymorphisms in VEGF (vascular endothelial growth factor) (Lambrechts et al., 2009), or H63D polymorphism in gene HFE (van Rheenen et al., 2013) and ALS have been reported. A literature search for peer-reviewed publications involving meta-analyses of the association between gene polymorphisms and ALS was conducted, and 26 articles with meta-analyses were identified (Chen et al., 2016: van Blitterswijk et al., 2012a: Lee et al., 2015: van Rheenen et al., 2013; Pan et al., 2016; Maiti et al., 2015; Pan et al., 2015; Lill et al., 2015; Zhao et al., 2016; Yang et al., 2015a,b; Chen et al., 2015; Smith et al., 2015; Keller et al., 2014; Ferrari et al., 2014; Diekstra et al., 2014; Govone et al., 2014; van der Zee et al., 2014; Wang et al., 2014a,b,c; Li et al., 2014; Fogh et al., 2014; Goris et al., 2014; ALSGEN Consortium et al., 2013; Ingram et al., 2012), but the question is that these conclusions drown in these studies are reproducible in larger studies.

3.1.3.3. ATXN2 gene. The results of a meta-analysis of observational studies showed that the odds of having intermediate degree CAG repeat expansion in the ATXN2 gene among ALS patients were about 4 times higher than those of the controls (Table 6 and Fig. 6) (Wang et al., 2014a). Comprehensive reviews and meta-analyses about this topic recently conducted by other investigators have led to similar conclusions (Neuenschwander et al., 2014; Laffita-Mesa et al., 2013).

3.1.3.4. Apolipoprotein E (APOE). The APOE  $\epsilon$ 4 allele may not be a risk factor for ALS (Govone et al., 2014; Praline et al., 2011; Zetterberg et al., 2008), unlike for Alzheimer's disease (Maiti et al., 2015). An association was reported between the  $\epsilon$ 4 allele and the occurrence of bulbar-onset ALS among men, but not women (Praline et al., 2011).

3.1.3.5. Early onset of balding (Androgenetic alopecia) and ALS. Alopecia is a highly heritable condition. A cohort study of 51,529 US men found that the individuals with extensive alopecia had about 3-fold increased risk (out of 42 diagnosed ALS cases, 18 reported moderate, 11 reported extensive alopecia) for ALS compared with those with no alopecia, and the risk of ALS increases linearly with severity of alopecia (p trend = 0.02) (Fondell et al., 2013). This observation may help interpret the male preponderance of ALS.

### 3.2. Risk factors associated with progression of ALS

One systematic review of risk factors related to the progression of ALS was identified (Chio et al., 2009a,b,c). The risk factors from this review have been summarized and updated based on publications that have since become available, focusing on modifiable risk factors. Details of the search strategy and summary can be found in Supplementary materials II and III. We also identified one new meta-analysis and four new observational studies for ALS progression in the update search (Supplementary material V).

## 3.2.1. Age at onset

Age at the onset of ALS was inversely related with survival time (Chio et al., 2009a,b,c). For every increase of 10 years in the age at onset, the HR (hazard ratio) for survival time was reduced by about 40% (Alonso et al., 2010a,b; Kamel et al., 2008; Pastula et al., 2009). For some early onset cases, the disease progression rate was slower, or even imperceptible for up to 40–50 years (Tosi et al., 1994). It is believed that these cases might be caused by environmental factors, some of which are discussed in this paper. Some environmental factors are modifiable and others are not, and the presence or absence of these factors affects disease progression (Doi et al., 2006).

### 3.2.2. Sex

Most studies have not observed sex effects on ALS progression (Chio et al., 2009a,b,c). However, higher mortality and shorter survival time in women with ALS was reported (Chio et al., 2009a,b,c; Kamel et al., 2008).

### 3.2.3. Ethnicity

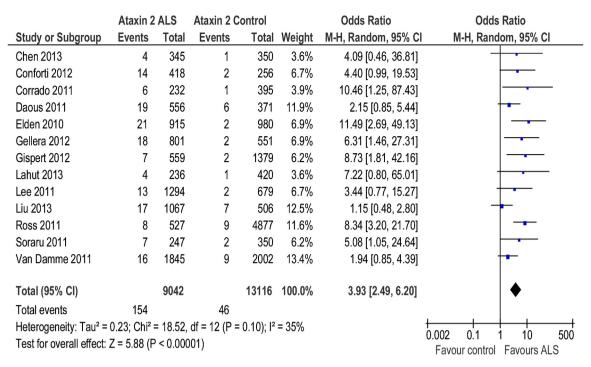
Ethnic Indians in Malaysia demonstrated a shorter interval between onset of symptoms and diagnosis, and shorter median survival time, when compared to non-Indians (Goh et al., 2011). However, analysis of Indians with ALS in Northern India revealed that among 1153 classic ALS patients, the overall median survival duration was  $114.8 \pm 25.9 (SE)$  months, with a range of 3-194 months, longer than their White counterparts (Nalini et al., 2008). Patients of North African origin were significantly younger, and had a shorter duration of disease, when compared to other ethnic groups (European, North African, Oriental, Balkan, Arab), and adjusted for age (Drory and Artmonov, 2007).

## 3.2.4. Onset site or ALS type

ALS bulbar onset is associated with a poorer prognosis than spinal onset. Compared to spinal onset, survival time among patients with bulbar onset has been estimated to be lower by about 40% (Pastula et al., 2009). In addition, the majority (about 60%) of young-adult patients showed pure upper motor neuron (*p*-UMN) signs confined to upper limbs, whereas among adult onset ALS cases, this proportion was less than 20%. Young-adult ALS with p-UMN phenotype had longer survival (74 months, 95% CI: 60.61-87.38) than did the classic phenotype (56 months, 95% CI: 48.65-63.34). In young-adult patients, a much higher proportion of men were in the p-UMN ALS group (5.8:1), whereas the ratio of men to women was 1.1:1 in the classic phenotype group (p=0.01)(Sabatelli et al., 2008). Among ALS phenotypes [classic, bulbar, flail arm, flail leg, pyramidal, respiratory, pure lower motor neuron (p-LMN) and p-UMN], the best outcomes were observed in p-UMN, pyramidal, p-LMN and flail arm phenotypes, and the worst in respiratory and bulbar phenotypes (Chio et al., 2011).

## 3.2.5. Progression rate

Disease progression rate is usually inversely associated with survival time. One good progression indicator is the interval



**Fig. 6.** The presence of intermediate CAG 30–33 repeats in the ATXN2 gene is associated with ALS. The data of intermediate CAG 30–33 repeats in ATXN2 were extracted from 13 included studies and the OR of the ratio of intermediate CAG repeat among ALS and control subjects was synthesized with meta-analysis using random effects model. Similar results were also obtained when using the fixed effects model (data not shown).

between the first sign of onset and diagnosis. Studies have shown that every year of increased time to diagnosis was associated with a nearly one year longer survival time (HR=0.77) (Pastula et al., 2009).

## 3.2.6. Body mass index (BMI)

BMI reduction rate from the BMI before disease onset to the time of the first visit is also a good indicator of disease progression. BMI reduction rate has been inversely associated with survival time among ALS patients (Kamel et al., 2008; Shimizu et al., 2012). The median BMI reduction rate was 2.5 BMI units per year (interquartile range: 1.3-3.8), and was significantly correlated with survival length (p < 0.0001) (Shimizu et al., 2012). A U-shaped relationship between BMI and mortality was reported, with the highest survival at 30-35 kg/m (O'Reilly et al., 2013). The adjusted HR for the linear association between BMI and survival was 0.860 (95% CI: 0.80-0.93, P=0.0001) (Paganoni et al., 2011). After adjustment, there was a 30% increase in risk of death for a 5% decrease from usual weight at time of diagnosis (RR = 1.30; 95% CI: 1.08–1.56). During follow-up, the adjusted mortality risk identified was increased by 34% (95% CI: 18-51%) for each 5% decrease in usual weight and increased by 24% (95% CI: 13-36%) for each unit decrease in usual BMI (p < 0.0001).

### 3.2.7. Nutritional status

Malnutrition during the disease course has been related to shorter survival (p=0.01), and fat mass associated with better outcomes (RR=0.90 for each  $2.5\,\mathrm{kg}$  of fat mass)(Marin et al., 2011b), high serum levels of triglyceride is significantly associated with a longer survival in ALS patients (Huang et al., 2015). A recent study also found that severe vitamin D deficiency accelerated the rate of decline by 4 fold, and was associated with a shorter life expectancy (Camu et al., 2013).

### 3.2.8. Comorbidity

Comorbidity would be expected adversely to affect the disease course of ALS. However, an early report showed that an abnormally elevated LDL/HDL ratio significantly increased survival time by more than 12 months for ALS patients (Dupuis et al., 2008). This was further supported by a univariate analysis showing that a higher LDL/HDL ratio was correlated with increased survival (HR = 0.9, p = 0.04), after adjusting for possible confounding such as age at onset, onset site and FVC (Sutedja et al., 2011). Comorbidities that were investigated in four subsequent studies included cardiovascular disease, dementia, Parkinson's disease, and depressive symptoms (Paganoni et al., 2011; Chio et al., 2009b; Korner et al., 2013; Dedic et al., 2012): none were found to significantly influence ALS disease course or the survival rate of patients (Korner et al., 2013).

## 3.2.9. Smoking

Smoking adversely affects survival time in patients with ALS. Compared with never smokers, the mortality RR (95% CI) was 1.45 (1.03–2.03) for heavy smokers (Alonso et al., 2010a,b). Current smoking was also independently associated with shorter survival time (HR = 1.51, 95% CI: 1.07, 2.15) (de Jong et al., 2012). In addition, the shortened survival time associated with smoking was more significant for women than for men. Compared with never smokers, the mortality RR (95% CI) for ever smokers was 1.31 (1.04–1.65) for women and 0.90 (0.72–1.11) for men, and the risk was particularly elevated for women who were heavy smokers (HR = 1.94, 95% CI: 1.24–3.06) (Alonso et al., 2010a,b).

## 3.2.10. Heavy metals

The survival interval from diagnosis to death showed a weak inverse association with blood lead level (HR = 0.9, 95% CI: 0.8–1.0) and a stronger inverse association with patella lead (0.5, 95% CI: 0.2–1.0) and tibia lead (0.3, 95% CI: 0.1–0.7) (Kamel et al., 2008). Of course, exposure to lead is not beneficial. These results might

indicate that the individuals with ALS caused by lead had longer survival times than did other types of ALS presentation.

## 3.2.11. ALS among veterans

The median survival time for veterans with ALS from symptom onset was 4.7 years (3.3 years from diagnosis). Veterans with ALS who had been deployed in Vietnam had significantly shortened survival time [HR = 1.73 (95% CI: 1.36–2.19)] (Pastula et al., 2009).

### 3.2.12. Health care

Improved survival time among ALS cases has also been reported over the last two decades in the USA, Japan and Italy (Testa et al., 2004; Czaplinski et al., 2006; Kihira et al., 2008), a trend that might be related to improved health care (Gil et al., 2009). A retrospective study from England showed that independent of any intervention, the median survival from diagnosis was 19 months, for patients who attended the multidisciplinary clinic, compared to 11 months, for those attending the general neurology clinic (HR = 0.51, 95% CI: 0.41–0.64) (Aridegbe et al., 2013). A retrospective study from France also showed that more aggressive use of non-invasive ventilation apparatus significantly improved survival from 2002 to 2009 (Gordon et al., 2012). However, the effect of health care improvement on ALS survival was not observed in Scottish data (Forbes et al., 2004).

### 3.2.13. Uric acid

A meta-analysis revealed that serum uric acid level was significantly (p < 0.0001) lower in ALS patients, particularly in cases with bulbar onset and longer disease duration (Abraham and Drory, 2014; Oh et al., 2015).

## 3.3. Genetic factors

## 3.3.1. ApoE

Five observational studies with the association between ApoE genotype and ALS progression could be identified. Two studies showed a significantly shorter survival time among ApoE  $\epsilon$ 4 carriers (Drory et al., 2001; Moulard et al., 1996). The third one showed reduced survival time among  $\epsilon$ 4 carriers with ALS, which was not statistically significant (Jawaid et al., 2011). The fourth and fifth studies showed no correlation between ApoE genotype and survival time (Lacomblez et al., 2002; Chio et al., 2016), but ApoE plasma levels were correlated with the rate of disease deterioration and survival time, with a relative risk of 0.647 (95% CI: 0.465–0.901; p=0.01) for decreased survival time (Lacomblez et al., 2002), and the presence of an ApoE  $\epsilon$ 2 allele significantly increased the risk of FTD (odds ratio, 2.61; 95% CI, 1.14–6.10; P=0.03) (Chio et al., 2016).

# 3.3.2. Other genetic factors potentially associated with ALS progression

The minor allele carrier status of rs12608932, a common variant located within an intron of *UNC13A* gene on chromosome 19p13.3, was strongly associated with approximately 1-year reduced survival of ALS patients (Chio et al., 2013a,b). A significant effect of 'long' polyalanine repeat alleles of NIPA1 was identified in a large European genomic study (van Es et al., 2009). In patients carrying 'long' alleles, median survival was 3 months shorter than patients with 'normal' genotypes, and the onset of symptoms occurred 3.6 years earlier. Common variants (rs9897526, rs34424835, and rs850713) and haplotypes of *GRN* (the gene encoding progranulin) were significantly associated with a higher mortality after onset of ALS (HR = 2.5) (Sleegers et al., 2008). Lower *SMN2* copy numbers and lower levels of estimated SMN protein (HR = 1.3, 95% CI: 1.1–1.6, p = 0.03) were associated with an increased mortality rate in ALS patients (Veldink et al., 2005a). Survival analysis showed that

there was a lower survival probability in those with the risk allele (C allele: CC and CT genotypes) of the *ZNF512B* gene (log-rank test, P < 0.01), independent of other prognostic factors in ALS (Tetsuka et al., 2013).

### 4. Discussion

ALS is a multifactorial neurodegenerative condition associated with modifiable environmental and unmodifiable genetic factors. Its rapid progression is mainly affected by un-modifiable risk factors, such age at onset and disease type. Therefore, the discussion below seeks primary to establish how likely it is for the risk factors for ALS onset identified in our analyses to be causal, as systematic review per se cannot be used to determine causality between risk factors and ALS.

### 4.1. Exposure to heavy metals

Exposure to heavy metals has been shown to increase the risk of ALS (Sutedja et al., 2009a,b; Johnson and Atchison, 2009; Ahmed and Wicklund, 2011; Vinceti et al., 2012; Iwami et al., 1994; Mitchell, 1987; Mitchell et al., 1991; Valentine et al., 2005; Trumbull and Beckman, 2009; Bowman et al., 2011; Caban-Holt et al., 2005). The most widely studied heavy metals in ALS patients are lead and mercury, both of which have been shown to be neuron toxicants in in vitro or in vivo studies, and to have accumulated in neural tissues (Kurlander and Patten, 1979). The modulation of normal SOD1 function by heavy metals is plausible because a causative link has been established between ALS and mutations in the SOD1 gene, from which the zinc/copper regulated superoxide degradation enzyme is synthesized. In addition, exposure to zinc either hastened or delayed the time of onset of ALS symptoms in an ALS animal model (Groeneveld et al., 2003; Ermilova et al., 2005). Further study is needed to clarify the role of zinc in the pathogenesis of ALS. A relationship between environmental selenium and ALS incidence has been reported (Vinceti et al., 2010). As the associations between ALS and zinc and ALS and selenium were largely inconclusive, the discussion will be focused on two heavy metals, lead and mercury.

### 4.1.1. Lead

A fuller discussion of the evidence on the association between lead and ALS has been published by (Wang et al., 2014b). Speculation about the association between exposure to lead and ALS began about 5 decades ago, when a series of ALS cases with antecedent lead exposure was reported (Oh et al., 2007; Livesley and Sissons, 1968). Since then, over a dozen retrospective casecontrol studies have been conducted. Most studies clearly showed that occupational exposure to lead was associated with a higher risk of ALS (Kamel et al., 2005; Scarpa et al., 1988), although some studies failed to observe a statistically significant association (Gresham et al., 1992; Pierce-Ruhland and Patten, 1981; Qureshi et al., 2006). The overall estimate of the risk of ALS effect due to previous exposure to lead in the meta-analysis conducted in this study was statistically significant. This association was supported by the following dose-response studies. One study showed an incidence gradient in the vicinity of a lead smelting factory in a county of the state of Missouri (Turabelidze et al., 2008). Kamel and his colleagues explored this association, with ALS cases and controls identified in Boston, using various lead measurement technologies (Fang et al., 2009; Kamel et al., 2002, 2005). A linear response relationship between lead exposure and ALS prevalence was confirmed in one case-control study (Fang et al., 2010), which was consistent with previous case-control studies (Fang et al., 2009; Kamel et al., 2002; Campbell et al., 1970). A major challenge in exploring the association between lead exposure and ALS is in the determination of historical exposure levels to lead. Since the lead levels in blood or other body liquids may not represent previous lead exposure, it is not surprising that some studies did not observe a difference in lead exposure between cases and controls (Kamel et al., 2005).

Paradoxically, the blood or bone levels of lead were positively associated with survival in a case-control study and in an animal study (Barbeito et al., 2010; Kamel et al., 2008). One possible explanation for this observation was that lead exposure might increase the production of antioxidants or stimulate VEGF (vascular endothelial growth factor) expression (Johnson and Atchison, 2009; Barbeito et al., 2010).

## 4.1.2. Mercury

Neurotoxic effects of mercury in humans and animals are well documented. In vitro experiments have shown that mercury damages the axons of neuron cells, a typical pathological change of ALS neuronal degeneration. Insufficient studies relating to mercury risk for ALS onset were identified for a meta-analysis. Although there are case reports of ALS associated with previous exposure to mercury, the relationship between the risk of ALS and exposure to mercury is yet to be established in analytical epidemiological studies. In vivo experiments have shown that mercury deposits in the nervous system and damages the axons of motor neurons, a pathological change that is typical of ALS neuronal degeneration (Pamphlett and Png, 1998). Chronic exposure to methyl-mercury induced early onset of hind limb weakness in a transgenic ALS mouse model, indicating that methyl-mercury promotes the progression of ALS (Callaghan et al., 2011), and may increase the risk of onset of ALS in humans.

### 4.2. Previous exposure to pesticides

An association between previous exposure to pesticides and ALS was identified in the systematic review and meta-analysis of observational studies conducted here, consistent with three recently publishedanalyses (Kamel et al., 2012; Malek et al., 2012; Kang et al., 2014), and reported similar risk estimates as ours. Evidence supporting the hypothesis that exposure to pesticides as a causal factor for ALS is discussed below.

## 4.2.1. Indirect evidence from genetic studies

Genetic studies have suggested that previous exposure to pesticides may be a strong risk factor for ALS (Morahan et al., 2007). The different variants of PON1 (paraoxonase/arylesterase 1), an enzyme which detoxifies organophosphate pesticides, have been associated with the occurrence of sporadic ALS in some publications (Saeed et al., 2006). A meta-analysis of the overall association between PON1 variant and risk of ALS was a minor increase [OR = 1.10 (1.01-1.20)] with borderline statistical significance (Wills et al., 2009), although the statistical significance was diminished if the included studies were adjusted for Hardy-Weinberg equilibrium (p > 0.05). However a causal association between PON1 and ALS should not be simply ruled out, because not all PON1 gene variants have been investigated. The sample size in each study might have been too small to demonstrate a significant association. A recent Italian study identified a higher frequency of PON1 gene mutations among ALS cases, compared to controls (Ticozzi et al., 2010); this article has provided the strongest evidence of an association between PON1 and ALS.

### 4.2.2. Chronic toxicity of organophosphorus compounds

A link between previous exposure to pesticides and ALS is supported by symptoms of chronic toxicity resulting from exposure to organophophorus compounds (OPCs). OPCs have been used as pesticides, insecticides, fertilizers, and nerve agents

in warfare for more than a century. OPCs can cause several types of neurological disorders in humans, depending on the exposure profile, amount, duration, and pathways (Jokanovic and Kosanovic, Organophosphate-induced delayed polyneuropathy 2010) (OPIDP) is thought to be similar to ALS. Investigators described an outbreak of OPIDP due to the contamination of soil by OPCs, which occurred in 1942 on a farm in Italy (Tosi et al., 1994). The patients in this outbreak were originally diagnosed with polyneuritis, with symptoms described as progressive, beginning with the limbs, but sensory nerves were not affected. The clinical manifestations would be consistent with current diagnostic criteria for ALS (Tosi et al., 1994). Similar syndromes have also been reported in some developing countries. A report from China described 143 cases with delayed dysneuria caused by pesticide poisoning, some of which resembled ALS (Zhang, 1991).

The most convincing epidemiological evidence of the association between exposure to pesticides and risk of ALS was reported in a prospective study, in which the mortality due to ALS in a pesticide factory was over 3-fold greater than expected (P < 0.05) (Burns et al., 2001). In addition, the increased incidence of mortality from ALS among soccer players, football players, baseball players, and among veterans could be partially explained by exposure to pesticides or other agricultural chemicals (Chio et al., 2009a; Vanacore et al., 2006; Belli and Vanacore, 2005; Wicks et al., 2007; Weisskopf et al., 2005a,b).

## 4.2.3. Evidence from case reports

ALS cases following exposure to agricultural chemicals have also been widely reported (Pall et al., 1987; Fonseca et al., 1993; Ahdab et al., 2011). The most relevant ALS case supporting a causal association between ALS and exposure to pesticides was reported (Doi et al., 2006). The disease progression in this case was modulated by the presence and absence of the pesticide.

Taken together, these observations favour the hypothesis that exposure to pesticides might be causally associated with ALS. This association could partially explain why professional soccer players (Chio et al., 2009a), foot players (Abel, 2007), and veterans who have ever participated in wars, but not professional basketball players or cyclists, have a higher risk of developing ALS.

### 4.3. Previous trauma

A history of trauma was identified as a risk factor from the first reported cases of ALS (Alpers and Farmer, 1949). It is important to determine how one or several repetitive local injuries could initiate ALS, a disease that progresses rapidly from the onset site to other parts of the body, for which hypotheses have been proposed (Gargiulo-Monachelli et al., 2012). The damage could spread, even it may spread further to the opposite motor cortex, the brainstem, and the spine. Two small-scale pathological studies supported the association between brain injury and ALS (McKee et al., 2009, 2013). The meta-analysis conducted in this study also identified a strong association between previous traumas and ALS. Further metaanalysis also revealed that ALS was associated with old trauma, which occurred at least 5 years prior to a diagnosis of ALS. This association supports the view that trauma may be a causal factor, because old trauma is unlikely to occur as a consequence of ALS. Therefore, although two recent studies argued that the traumas were the consequence instead of the cause of ALS (Beghi et al., 2010; Turner et al., 2010), previous trauma should still be considered a potential risk factor for ALS (Pupillo et al., 2012).

If previous trauma is a risk factor for ALS, then some risk factors associated with ALS – professional sports, lower BMI, lower educations, and strenuous work – may have been confounded by previous trauma. All of these potential associations are discussed further below.

## 4.3.1. Physical activity and participation in organized sports

Participating in sports or physical activities has been associated with ALS in some, but not all, early studies. If professional sports were excluded, the overall risk estimate from our meta-analysis showed that participating in sports was not significantly associated with an increased risk of ALS. This result is consistent with a comprehensive study from Veldink's group (Veldink et al., 2005b). However, we did find that participating in organized or professional sports was significantly associated with ALS. These different estimates could explain some conflicting observations in literature about the association between sports and ALS (Longstreth et al., 1998; Veldink et al., 2005b; Qureshi et al., 2006; Valenti et al., 2005). Many studies did not distinguish general physical activity and sports from the professional organized sports (Okamoto et al., 2009b), and summarized the results together (Hamidou et al., 2014). The possible explanations for the difference between professional and recreational sports are that professional athletes have a higher risk of getting trauma/injury (Harwood et al., 2009), and exposing to pesticides (Su et al., 2016).

## 4.3.2. Blue collar occupation and lower education

As individuals involved in blue collar occupations often possess lower education levels, thus these two risk factors are correlated. Early epidemiological studies, especially those conducted in Italy, found that workers having strenuous occupations had a higher risk of developing ALS (Chancellor et al., 1993; Gunnarsson et al., 1991; Buckley et al., 1983). These occupations required repetitive motions or heavy labour, such as professional sports, occupations related to primary and secondary industries, including construction and mining. Repetitive work itself may not be able to cause motor neuron degeneration which occurs in the brain motor cortex, the spine, and the brainstem through common pathological change. Many potential risk factors, such as exposure to pesticides, mechanical injury, exposure to heavy metals, and smoking have been associated with ALS, but actually these potential risk factors are highly correlated with blue collar occupations and could be good surrogate variables for blue collar occupations.

## 4.3.3. High fitness level

Three publications reported an association between having a high fitness level and ALS. The first article showed an inverse association between BMI and the incidence of ALS among college varsity athletes (Scarmeas et al., 2002). Two subsequent studies confirmed this observation, showing that the incidence of ALS was less common in patients with cardiovascular disease (CVD) than in non-CVD patients (O'Reilly et al., 2013; Sutedja et al., 2011). These two studies concluded that comparing to control population, the BMI value was significantly lower in ALS patients prior to the onset of disease, along with significant higher physical strength. However, these three studies could not rule out confounding effects from head and spinal trauma/injuries. The individuals with lower BMI in these studies were army or professional sport participants. They are physically active, and were more likely to be injured than individuals with higher BMI in general population. In addition, a higher BMI has been consistently demonstrated being a protective factor against ALS progression (Paganoni et al., 2011), the actual involved mechanism is unknown.

## 4.4. Lifestyle

## 4.4.1. Smoking

Early studies produced controversial results about the association between smoking and ALS (Kamel et al., 1999; Fang et al., 2006; Qureshi et al., 2006; Nelson et al., 2000). In the present

analysis, we found that the association between smoking and ALS was significant in females, not in males. Since a high proportion of bulbar onset ALS cases have occurred in females, smoking may be a risk factor for bulbar ALS, or may promote early onset in females (Korczyn and Drory, 2010). A meta-analysis of evidence from cohort studies suggested that smoking was a causal risk factor for ALS in females, with former smokers found to be at higher risk of ALS (Wang et al., 2011a,b). Since males tend to be more frequently exposed to other potential ALS risk factors, such as pesticides or organic solvents during their working life, the lack of associations in males might have been confounded by occupation. More population-based studies designed to study ALS aetiology are required, as most of the studies quoted in this project were not designed specifically for that (Wang et al., 2011a,b; Gallo et al., 2009). Even though smoking is not a strong risk factor for ALS, patients should be encouraged to quit smoking because smoking has been associated with higher incidences of other chronic diseases, including cardiovascular disease and cancer. One potential mechanism by which smoking could affect the risk of ALS is the possible impairment of ER (endoplasmic reticulum) and mitochondria (Federico et al., 2012) by the oxidization products from smoking (Alonso et al., 2010a,b). The contamination of tobacco by heavy metals and pesticides could be another possible explanation for the association between ALS and smoking (Weisskopf et al., 2010).

## 4.4.2. Alcohol and coffee consumption

Light to moderate alcohol consumption has been reported to be inversely associated with the incidence of dementia and Alzheimer's disease (Daviglus et al., 2011). An observational study recently found that current alcohol consumption was associated with a reduced risk of ALS (de Jong et al., 2012). Our preliminary meta-analysis does not support an association between alcohol consumption and ALS risk. However, this association might not indicate a protective effect of alcohol *per se*, because a risk factor must exert its effect prior to ALS onset, rather than contemporaneously. The same argument also could apply to the inverse association between coffee consumption and ALS (Beghi et al., 2011).

### 4.4.3. Vitamin E intake

It has been suggested that oxidative stress resulting from reactive oxygen species (ROS) is associated with ALS onset. In comparison to controls, oxidative products in plasma and SOD activity in erythrocytes were elevated in ALS patients (Bonnefont-Rousselot et al., 2000; Oteiza et al., 1997). Increase in lipid peroxidation in spinal cord was observed prior to the onset of ultrastructure or clinical changes in SOD1 transgenic mouse model of familial ALS (Hall et al., 1998). The greatest intensity of motor neuronal injury due to lipid peroxidation was observed in ALS patients in the active phase of disease progression (Hall et al., 1998). Oxidant treatment of cultured motor neurons has caused a dose-dependent increase in apoptosis (Kaal et al., 1998). In addition, an epidemiological case-control study found that a higher intake of food rich in antioxidants, such as fruit and vegetables, might play a protective role against the development of ALS (Okamoto et al., 2009a). Therefore, intake of vitamin E, an antioxidant, could be a protective factor against ALS. However, supplementary vitamin E given to ALS patients was ineffective in arresting disease progression. After 12 months of treatment, supplementary vitamin E had no effect on disease progression or survival time (Desnuelle et al., 2001; Graf et al., 2005; Galbussera et al., 2006), indicating that progression of the disease was not reversible by vitamin E.

#### 4.5. Electric shock

Early epidemiological studies showed that occupational exposure to electromagnetic fields was positively associated with ALS (Li and Sung, 2003), and a synthetic OR, based on a meta-analysis, was significantly elevated (Zhou et al., 2012). However, animal experiments failed to establish an association (Johnson and Atchison, 2009; Poulletier de Gannes et al., 2009), raising the possibility of confounding by other risk factors in some human population studies. A large prospective study in Switzerland found that distance from the magnetic field resulting from hydro powerlines was not associated with an altered risk of developing ALS (Huss et al., 2009). Since our meta-analysis showed an association of ALS with a history of electric shock, the association of occupational exposure to electromagnetic fields with ALS could be confounded by a history of electric shock. Many ALS cases have been diagnosed following electric shock (Al-Ajmi et al., 2012; Sirdofsky et al., 1991; Rose, 1994). These case reports form the strongest evidence of a causative association between electric shock and ALS. A review indicated that the evidence linking electrical occupations to an increased risk for ALS was remarkably consistent, but that the evidence of an association with measured magnetic field levels was weak (Consales et al., 2012). In agreement with our conclusion, a recent meta-analysis of observational studies did not support an association between extremely low frequency magnetic fields and ALS (Vergara et al., 2013). Based on these observations, we hypothesize that electric shock is the one of confounding factors, although the mechanisms by which electric shock could cause ALS are not known, and a low quality systematic review failed to reveal any association with ALS (Abhinav et al., 2007).

### 4.6. Intake of beta-N-methylamino-L-alanine (BMAA)

The incidence of ALS-PDC (ALS-Parkinson-Dementia-Complex) syndrome, in Guam and surrounding Pacific islands, was found to be up to 50-100 fold greater than the average incidence of ALS worldwide during the 1940s to the 1960s. It did decline to rates similar to those reported in other regions after 1980, before its cause was established (Stone, 1993). The excess cases of ALS-PDC in Guam and surrounding Pacific islands were first associated with the food toxicant BMAA in 1987, after partial replication of both clinical and pathological changes typical of ALS-PDC resulting from high doses of BMAA in monkeys (Spencer et al., 1987a,b; Spencer, 1987). BMAA is a neurotoxic amino acid, which can be produced by most aquatic cyanobacteria. BMAA neurotoxicity occurs via the activation of glutamate metabotropic receptors (mGluR5 and mGluR1) and the induction of oxidative stress (Banack et al., 2010). This hypothesis was initially overlooked, possibly due to negative follow-up studies in mice (Perry et al., 1989; Cruz-Aguado et al., 2006), and the discovery of less than sufficient residual amounts of BMAA in the cycad seed flour, which was consumed by the local residents in Guam, after conventional washing processes (Duncan et al., 1990). However, this unproven hypothesis was revived 15 years later by Cox et al. (2009), after bio-magnification of BMAA was observed in foxes, which used to be consumed by the Chamorro population. The presence of BMAA in brain tissue was also reported in ALS-PDC cases (Murch et al., 2004). Clinical and pathological changes mimicking ALS in rats has resulted from exposure to BMAA (de Munck et al., 2013). Despite the return of the BMAA hypothesis, causal association has still not been established in humans. The disease term ALS-PDC was excluded from our selection of ALS disease terms since we were not sure that the form of ALS seen in the Chamorro people was same as sporadic ALS in other regions of the world. It was believed that the environmental risk factor associated with ALS-PDC was unique to Guam and the surrounding areas when this work was initiated. Since the present study was designed to synthesize epidemiological studies, most studies related to BMAA would have been excluded even if ALS-PDC term was included in the search.

The high incidence of ALS in the Pacific islands has declined over the last four decades, which may be due to lifestyle and dietary changes, and is now close to the incidence reported in other regions. Whether this food toxicant theory is true or not, the story of the temporary ALS epidemic in Guam and other Pacific islands unequivocally indicates that environmental factors are associated with ALS. Similar temporal fluctuations in ALS incidence, but to a much lower degree, has also been observed in many other areas worldwide (Gordon et al., 2011a; Doi et al., 2010; Alonso et al., 2011; Georgoulopoulou et al., 2011), further underscoring the role of environmental factors in the pathogenesis of sporadic ALS.

## 4.7. Military service

Military service may not be an independent risk factor for ALS because the number of wars participated is linearly associated with ALS. The organic phosphorus compounds present in explosives and war associated trauma/injuries (Bergman et al., 2015) or exposure to chemicals (Beard and Kamel, 2015) might be the underlying risk factors for ALS in military personnel or veterans.

## 4.8. Viral infection

Because ALS specifically affects voluntary skeletal muscles, an immune response elicited by viral infection might cause paralysis in a tissue (motor neuron) specific manner, leading to the speculation that viral infection may be a risk factor for ALS. Such speculation is consistent with the abnormal immune response observed in ALS patients (Rentzos et al., 2011). The hypothesis of viral infection and subsequent immune response could explain the specificity of ALS affected tissues. Unfortunately, convincing epidemiological evidence supporting this association has not been found.

# ${\bf 5.}$ Possible biological interpretations of risk factors associated with ALS

Risk factors for development and progression of ALS may be considered with the context of the process of prion-like propagated protein misfolding, which is relevant not only to ALS, but to other neurodegenerative diseases such as Alzheimer's and Parkinson's disease (Guest et al., 2011). Two proteins participate in this prion-like propagation of protein misfolding in ALS: CuZn superoxide dismutase (SOD1) (Grad et al., 2011; Smethurst et al., 2014) and TDP43 (Nonaka et al., 2013). Both proteins are known to become misfolded and aggregated under various conditions of cell stress. Ataxin-2 may participate in the expression of 'seeding' proteins or co-factors in their generation (Wang et al., 2014a). Exposure to heavy metals such as lead and mercury, as well as to pesticides and solvents, is known to posttranslationally modify or oxidize proteins in vivo (Stohs and Bagchi, 1995; Braconi et al., 2011) and could be associated with formation of a productive nidus or 'seed' which can propagate its misfold intracellularly and intercellularly. Smoking-associated formaldehyde exposure can also post-translationally modify/cross-link lysine, arginine and tryptophan in proteins (Edrissi et al., 2013; Reisner and Lundblad, 2009), perhaps including those proteins that can provide a productive seeding template for ALS. Electric shock could heat-denature proteins (Tsong and Su, 1999) to potentially form a productive misfolded protein seed that could propagate to non-injured regions. Military service may has also been thought to be associated with exposure to BMAA (Cox et al., 2009), which can disturb conformation of proteins though incorporation of BMAA as a natural but non-protein amino acid (Glover et al., 2014), additional to the association with trauma and explosive compounds. Heavy work and strenuous athletic activity are associated with the generation of reactive oxygen species (Parker et al., 2014), which may change protein conformation via oxidative modification of the free-radical defence enzyme SOD1. Viral infection can alter protein translation and chaperone availability of proteins in the nervous system. Once a propagated protein misfolding process has been seeded, different factors may impact on the rate of progression of disease, including the availability of fat reserves to tolerate decrease nutritional intake, vitamin D deficiency, comorbidities, and other factors that are not related to the initial genesis of the original template for misfolding.

### 6. Conclusions

ALS is a multifactorial neurodegenerative condition that affects primarily voluntary muscles. No single risk factor has been identified that explains a significant proportion of total ALS cases. Mutations in the two most common monogenic risk factors, SOD1 and C9ORF72, could explain about 25% and 35%, respectively, of familial ALS cases. All the monogenic mutations, in over 30 genes, could explain about 10% all ALS cases (Chio et al., 2012). For all environmental or other genetic risk factors identified in the present study, each risk factor was associated with only a small proportion of total ALS cases.

One genetic factor (ATXN2) and five environmental factors (previous exposure to heavy metals including lead, previous exposure to pesticides, a history of physical trauma/injury, a history of electronic shock, and previous exposure to organic solvents) have been identified with relatively strong associations with the onset of ALS in this study, supporting multi-factorial and multi-pathway etiology of ALS. Further studies are required to verify these findings. With respect to the progression of ALS, age and disease type were the two most important risk factors determining the survival of ALS patients. Modifiable risk factors, such as nutrition status and BMI level, have also been found to modify the survival time of individuals with ALS. Genetic risk factors may also modify the progression of ALS. Although the current study has identified a number of potential risk factors associated with the onset or progression of ALS, further work is required to establish causal associations. In addition, as the number of studies included in most meta-analyses in this study was small, caution should be used when interpreting the findings.

## Conflict of interest

The authors declare that there are no conflicts of interest.

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### Appendix A. Supplementary data

Supplementary data associated with this article can be found, in the online version, at <a href="http://dx.doi.org/10.1016/j.neuro.2016.06.015">http://dx.doi.org/10.1016/j.neuro.2016.06.015</a>.

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